# Oligopolies in Trade and Transportation: Implications for the Gains from Trade

Enrico Cristoforoni[*]      Marco Errico[†]      Federico Rodari[‡]      Edoardo Tolva[§¶]

**Preliminary. Do not circulate.**

## Abstract

We study the interplay between oligopoly in the transportation industry and oligopsony power retained by non-atomistic importers. We leverage transaction-level data from Chilean customs to document several empirical facts: (i) strong concentration in the transportation sector and among importers, and (ii) that transportation prices are highly dispersed and are the outcome of bilateral negotiations in the presence of dual market power. We develop and estimate a trade model that departs from the usual iceberg cost assumption and allows for two-sided market power in the transportation industry. We find that transport carriers charge large markups, but importers benefit from substantial bargaining power. Finally, we embed the bilateral bargaining framework into a quantitative trade model of importing. We show that the welfare costs of higher tariffs are 40% lower because of the presence of importers' market power. We also show that the pass-through on aggregate welfare of carbon policies on the shipping market (such as the EU ETS) have negligible effects.

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[*]Boston College. Contact: cristofe@bc.edu
[†]Bank of Italy. Contact: marco.errico@esterni.bancaditalia.it
[‡]Boston College. Contact: federico.rodari@bc.edu
[§]University of Warwick. Contact: edoardo.tolva@warwick.ac.uk.

# 1 Introduction

Every year, over $20 trillion of international trade flows are carried around the globe by the transportation sector. Recent events such as the 2021 Suez Canal obstruction, the severe port congestion in 2021-2022, and the piracy and terrorist attacks in the Red Sea in 2023 and 2024 have demonstrated the crucial and indispensable role of the transportation sector for global trade and economic growth. Yet, little is known about the market structure of the transportation sector and how prices in the market for transportation services are determined.

In this paper, we study how imperfect competition and bilateral negotiations in the international transportation sector impact the determination of transportation prices, and their implications for trade flows, gains from trade, and shock transmission. Most conventional quantitative trade frameworks assume the presence of iceberg trade costs, treating the price of transportation services as an exogenous multiplicative friction. However, we document that the transportation sector is highly concentrated, and transportation prices are the outcome of bilateral negotiations between transportation carriers and importers. These features are usually absent from models in the trade literature.

Leveraging detailed Chilean import customs data from 2007 to 2022, we document several key empirical facts about the market structure of the transportation sector and the determination of transportation service prices. For each international shipment, customs authorities collect detailed information on freight costs, the transportation mode, and the carrier performing the last leg of the shipment before clearing customs, along with standard information such as importer identity, product code (HS8), country of production and origin of the shipment, and the value and quantity shipped. These features of the data allow us to measure freight unit prices at the shipment level.

The customs transactions enable us to document a high degree of concentration in the transportation sector. We document that, from the same origin, multiple modes of transportation are used, but importers tend to rely on a single mode. However, each importers uses multiple carriers, and carriers serve multiple firms at the same time. This motivates us to define a market in which carriers compete as a combination of mode of transportation (sea, air, and road), country of origin, and HS4 sector. We then construct the Herfindahl-Hirschman Index (HHI) by computing the share of imports that each carrier operates within a given market. The average HHI index across markets is 0.55, far above the threshold value of 0.25 often used to define an industry as oligopolistic. This evidence suggests that carriers exerting some degree of market power might be relevant in international shipping.

We provide reduced-form evidence in line with carriers and importers negotiating bilaterally over the price of transportation services. The data reveal significant dispersion in the unit freight price that a carrier charges within the same market. We show that the coefficient of variation in unit freight cost within a carrier-market is, on average, 0.9 and not proportional to the value of the shipment, in contrast to the "iceberg" assumption. A statistical decomposition of unit freight price dispersion within a carrier-market-time combination further shows that 89% of the variation is specific to the carrier-importer relationship. Lastly,

within a carrier-market pair, unit freight prices decrease with the importer's share of the carrier's total quantity and increase with the carrier's share of the importer's total imports. This is consistent with the presence of importer-carrier bilateral bargaining and buyer and seller market power.

Inspired by these facts, we develop a theory in which transportation carriers and importers bilaterally bargain over the price of transportation services, exerting both seller and buyer market power. Carriers exert seller market power as they compete in a standard oligopoly due to their non-atomistic nature, in line with the empirical evidence. Importers are also non-atomistic and exert buyer market power (oligopsony) as they internalize the upward-sloping supply curve of the carriers. We assume carriers and importers undertake Nash-in-Nash negotiations, and equilibrium unit freight prices reflect the relative bargaining power along with oligopoly and oligopsony incentives.

We bring the theory to the data to structurally estimate the relative bargaining power of carriers and importers. Following Alviarez et al. (2023), we leverage the network structure of the data to identify and estimate the key parameters via GMM. Using Hausman-type instrumental variables, we first estimate the within-market substitutability across carriers to be 3, indicating the presence of sizeable markups in the transportation sector. On average across markets, importers exert buyer market power due to a carriers' supply elasticity of 0.43, and they are estimated to have 2.5 times more bargaining power than carriers. We then show that both buyer market power and importers' bargaining power correlate positively (negatively) with the number of carriers (importers) operating in the market, validating our estimates.

We embed the bilateral bargaining framework into a quantitative trade model of importing to assess the effects of imperfect competition and bilateral negotiations in the transportation sector. The economy is populated by a finite number of heterogeneous domestic firms that produce and sell a differentiated product both domestically and internationally, in the presence of roundabout production (Caliendo and Parro, 2015). Firms have the option of importing a bundle of foreign intermediate inputs, subject to fixed costs of importing. Imported inputs increase firm productivity as they imperfectly substitute domestic inputs (Halpern et al., 2015; Blaum et al., 2018). Importers engage in Nash-in-Nash negotiations over the unit freight price in a transportation market populated by a finite number of heterogeneous carriers with an upward-sloping supply curve.

We quantify the model using a combination of customs data and firm-level balance sheet information for the manufacturing sector. We estimate five parameters using a two-stage Simulated Method of Moments. We discipline the distribution of domestic market shares, carriers market shares, and bilateral market shares targeting key moments of the cross-sectional distributions of market shares from the data. We use firm-level balance sheet information from the survey of manufacturing industries (ENIA) to calibrate the key parameters of the domestic production process such as final producers' demand elasticity and the substitutability between domestic and imported inputs. The estimation process yields an overall good fit in matching targeted moments.

Using the estimated model, we show that the increase in consumers' prices due to tariffs is 40% lower in our model compared to a standard model with iceberg trade frictions thanks to the presence of decreasing return to scale and importers' market power. The lower welfare costs of tariffs in the presence of endogenous trade costs is driven by a decline in trade costs, that partially offset the rise in the factory-gate price of imported goods. Moreover, entry/exit of importers further lower the costs of tariffs for a sufficiently low ad valorem tariff. The reason is that surviving importers can benefit from stronger buyer market power vis-a-vis carriers, which allows them to further reduce their cost of transportation and further offset the rise in the price of imports due to the tariff.

We show that carbon policies such as the extension of the EU ETS on the shipping market have negligible effects on aggregate welfare, but the asymmetry of the policy generates strong reallocation across carriers. The quantitative exercise applies a €50 tax per ton of $CO_2$, which is mapped to be equivalent to a 5 p.p. increase in carriers' marginal cost. The welfare cost for final consumers is negligible because the pass-through of the carbon tax into transportation prices is incomplete and, more importantly, transportation costs account for a small share of importers' marginal cost. In a scenario in which the policy is applied to only a subset of carriers, all carriers increase their price. While the carrier facing the carbon tax is forced to increase the price, the other carriers benefit from increasing their seller market shares and their markup on importers.

**Related Literature**    We contribute to the large literature in international trade focusing on the role of trade costs (Anderson and Van Wincoop, 2003; Eaton and Kortum, 2002). We contribute to this literature considering the features and the market structure of the transportation industry, endogeneizing transportation prices in a model of bilateral bargaining with both buyer and seller market power. Early work by Hummels et al. (2009) explores the role of market power and price discrimination in the shipping industry using aggregate data. More recently, with the availability of transaction level data,[1] Ignatenko (2020) shows firm size affects freight charges in a model of price discrimination through quantity discounts; Asturias (2020) explores how the number of shipping firms impacts transport prices and trade flows. In a model with market power in the upstream transportation sector, Ardelean and Lugovskyy (2023) introduce search and information frictions in the downstream importing sector. Brancaccio et al. (2020) also endogeneize transport costs in the presence of search frictions between exporters and transport firms. Wong (2022) endogeneizes transportation prices in a model with "round-trip" effect.[2]

Our paper contributes also to the limited literature on firm-to-firm trade. Novel empirical work is made possible by the recent availability of domestic and international firm-to-firm transaction data. Using domestic firm-to-firm transactions, Alfaro-Urena et al. (2022) focus

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[1]Ardelean et al. (2022) surveys the recent academic research on maritime shipping stemmed from the explosion of new micro data sets in shipping and international trade.

[2]Our work is also tangentially related to the literature on port infrastructure, network effects, and congestion (Heiland et al., 2019; Do et al., 2024; Ganapati et al., 2021).

on the effects of establishing a first linkage with a multinational buyer; Dhyne et al. (2022) develop a model of oligopolistic competition in firm-to-firm trade to rationalise the positive relationship between suppliers' markups and the supplier's share among their buyers. Using international firm-to-firm transactions, Alviarez et al. (2023) develop a pricing theory of firm-to-firm trade accounting for both oligopoly and oligopsony forces in the U.S. import data; Gopinath and Itskhoki (2011) and Grossman and Helpman (2020) develop theories of bargaining in firm-to-firm trade across borders and discuss the implications for exchange-rate pass-through and the organization of supply chains, respectively. Our contribution is new empirical evidence from the international transportation market and the focus on the implications for trade flows and the gains from trade.

Our quantitative analysis and counterfactual exercises contribute to the literature that measures how consumer welfare is affected by international trade (Arkolakis et al., 2012). Our quantitative framework is inspired by models of input trade with firm heterogeneity, e.g. Halpern et al. (2015) and Blaum et al. (2018). Relative to these works, we quantify the welfare impact in the presence of endogenous trade costs arising from bilateral bargaining. Our work is also related to the literature investigating the determinants of shock transmission across borders. Several works have shown the importance of market structure for the pass-through rate of exchange rate fluctuations (Atkeson and Burstein, 2008; Amiti et al., 2019a) or tariffs (Alviarez et al., 2023; Amiti et al., 2019b). In this regard, we leverage our quantitative framework to study how endogenous transportation costs influence the transmission of tariff shocks.

The rest of the paper is structured as follows: Section 2 provides a set of stylized facts on the transportation industry and bilateral bargaining on unit freight prices. Section 3 presents and structurally estimates the model of bargaining over unit freight prices. Section 4 describes the quantitative model, its estimation, and the counterfactual exercises. Section 5 concludes. The Appendices contain additional tables and figures, derivations of key theoretical results, and additional data and estimation details.

## 2    Stylized Facts on Dual-Market Power

### 2.1    Data

We use transaction-level data on imports from Chilean Customs covering the period 2007-2022. The data includes information on the party conducting the transaction such as the importer, the exporter, the product, the mode of transport (sea, air, and road freight), the port of entry [3], and the country of origin. There is also information on the content of the transactions themselves such as the product code (HS8), the weight, the number of items, both the CIF and FOB values, and freight values. More importantly for this project, we also

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[3]The Chilean custom refers as port of entry where custom declaration for the good is processed and can be either a maritime port, an airport or a border entry point.

observe the name of the shipping company that took care of the transportation of the goods. A full list of the variables used can be found in Appendix C.

We collapse the data at the yearly level by importer-country of origin-carrier-product-transport mode. A key challenge in the data cleaning process is the identification of the carrier company. The data are not standardized, and the name of the carrier company is often misspelled or written in different ways. We use a combination of string matching and manual inspection to identify the carrier company. More details about the data cleaning process can be found in Appendix C.2.

We choose to focus our attention on imports for two reasons. First, Chilean exports are mostly driven by trade in commodities, with ores and metals accounting for almost 50% of the total value.[4] Second, and more important to our analysis, customs data contain information on the party responsible for arranging the shipping contract, the so-called INCOTERMS - the International Chamber of Commerce's International Commerce Terms. According to INCOTERMS, any transaction is classified in two categories depending on whether it is the importer's or the exporter's responsibility to arrange the international shipping of the good. In line with previous literature (Ardelean and Lugovskyy, 2023; Teshome, 2018), we find that the majority of international shipments are arranged by the importing firm (Figure E.1 in Appendix E).

**Structure of Chilean Freight Market** In line with aggregate statistics on international trade and international shipping [5], Figure E.1 in Appendix E shows that also in our sample, maritime transport is the most used mode of transport, with more than 50% of transactions conducted by sea. Air transport is the second most used mode, accounting for around 40% of the transactions in the sample. Road transport is seldom used given the geographical distance between Chile and its main trading partners. However, in terms of the total volume of trade, maritime transport is predominant both in terms of value and weight. The significant discrepancy between the share of transactions and the share of value and weight is attributable to the less frequent use of maritime transport compared to air transport.

Figure E.2 in Appendix E illustrates that firms tend to use a single mode of transport for the majority of their transactions.[6] Approximately 80% of importers use a single mode of transport for each origin-product pair. However, multiple modes are used for imports from specific countries. This pattern suggests that the choice of transport mode is not solely determined by a combination of the country of origin and the characteristics of the imported goods. This observation motivates our definition of a market as an origin-HS4-mode triplet.[7]

Despite using few modes of transportation, importers interact with multiple transportation companies. Table E.3 in Appendix E classifies all carrier-to-importer matches into four

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[4]Source UNCTAD statistics on Chile year 2022.

[5]Refer to Ardelean et al. (2022) for a recent survey.

[6]Appendix E.1.3 provides information on sectoral and sourcing composition of Chilean imports, both at the aggregate and at firm level (Figure E.2). Chilean firms tend to import from a limited number of countries (Figure E.3), in line with broad evidence from international trade.

[7]In figure E.4 in Appendix E we show that the median firm trade only with 2 product (HS4) in the sample.

groups: one carrier to one importer, one carrier to multiple importers, multiple carriers to one importer, and multiple carriers to multiple importers. We show that both importers and carriers interact with other firms in most of the linkages, as the share of many-to-many imports is almost 60%. The remaining fraction of imports and linkages is classified as one-to-many, in which one carrier has relationships with many importers. Not surprisingly, one-to-one and many-to-one trades are marginal. These features of the network in the transportation market highlight how bilateral bargaining can play a key role in shaping the market equilibrium.

## 2.2 Stylized Facts on Market Power in Trade and Transportation

In the following section, we use transaction-level data to show that both imports and freight carriers are highly concentrated, and that transportation costs exhibit evidence consistent with bilateral bargaining and two-sided market power.

**Concentration in Trade and Transportation**    Figure 1 shows that imports and freight carriers are characterized by the presence of large firms. The left panel shows that only a handful of firms drive aggregate imports, in line with previous literature (Bernard et al., 2007; Mayer and Ottaviano, 2008). The red curve plots the cumulative distribution of imports in 2015 after ranking importers from the left to the right, starting with the biggest. The top 0.1%, 1%, and 10% of importers account for 35%, 65%, and 95% of total imports, respectively. The right panel reports the average HHI index across different markets over time, showing that concentration is high across freight carriers. We construct the HHI index by computing the share of imports in value that each carrier ships within a given market. We define a market as a combination of mode of transportation (sea, air, and road) and country of origin or, alternatively, conditioning on the mode, country of origin, and HS4 sector, as we believe key competitive forces operate within routes (and are potentially product-specific). Using both definitions of market, the average HHI across markets is above the threshold value of 0.25 (approximately 0.3 and 0.55, respectively), indicating the presence of strong concentration among freight carriers.[8]

As robustness, Figure E.2 in Appendix E shows that concentration among carriers exhibits the same quantitative dynamics if market shares are measured in terms of quantity or weight (in kilograms) shipped. In addition, Figure E.1 plots the entire distribution of HHI indices when a market is defined as a mode, country of origin, and HS4 sector combination. Most of the markets exhibit moderate or high concentration, with indices above the 0.15 and 0.25 thresholds, with no differences between modes of transportation.

**Variation in Bilateral Freight Prices**    We show that freight prices vary substantially within market and within carrier, and the importer-carrier match-specific component explains a sub-

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[8]We also consider more aggregate definition of markets, such as the aggregate set of imported goods or just conditioning on the mode of transportation. In these cases, Figure 1 shows that HHI indices are lower than the standard threshold used in defining a market as competitive. Moreover, despite multiple mergers and acquisition in the shipping industry, Figure 1 shows that concentration has not increased over the last 15 years.

Figure 1: Concentration among Importers and Freight Carriers
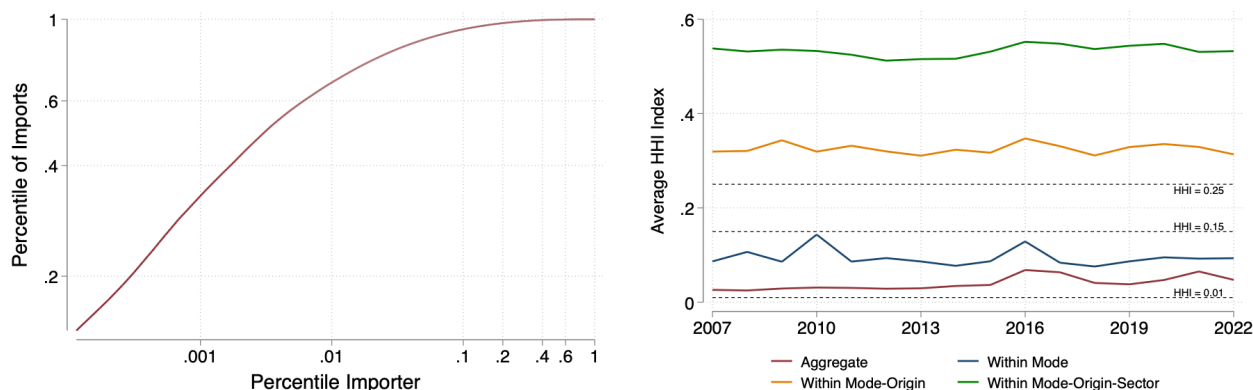


**Notes**: The left panel plots the cumulative distribution of importers for the year 2015. Importers are ranked according to their size from left to right on the horizontal axis. The vertical axis reports the cumulative contribution to aggregate imports. Axis are in log scale. The right panel plots the average HHI index across the different markets of the transportation sector over time. Markets are defined according different levels of granularity. The red line considers a unique aggregate transportation market. The blue line defines markets by their mode of transportation (i.e. sea vs air vs road freight). The orange and green lines defines markets as a combination of mode-origin and mode-origin-sector, respectively. A sector is defined as a HS4 category. Carriers' market share are computed in terms of value shipped.

stantial portion of the variation in freight prices.

Figure 2 shows that unit freight prices are highly dispersed even within carriers, contrary to widespread modeling assumptions. For each market (sector-origin-mode-time combination), we compute the coefficient of variation (CoV) of unit freight prices, computed by dividing total freight cost by the quantity transported. The mean and median CoV across markets is approximately 0.9 and 0.8, respectively, indicating the presence of substantial dispersion in prices (Ignatenko, 2020; Ardelean and Lugovskyy, 2023).[9] We also find that carriers within the same market discriminate across importers, charging different unit freight prices, as most of the dispersion survives after conditioning also on carriers.

As robustness, Figure E.3 in Appendix E shows that the dispersion in unit freight prices is quantitatively similar when we measure unit freight prices in terms of unit of value shipped or per kilogram. Similarly, the distribution of coefficients of variation is similar across modes of transportation, suggesting that price discrimination is quantitatively similar in sea and air freight, and slightly lower in road freight. Lastly, Figure E.4 in Appendix E shows that unit freight prices are not directly proportional to the shipment value, indicating that the data reject the standard iceberg trade cost assumption.

Table 1 further shows that most of the dispersion in unit freight prices is explained by a carrier-importer match-specific component, indicating the presence of bilateral forces in determining freight prices. We follow Fontaine et al. (2020) and consider the following

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[9]Fontaine et al. (2020) and DellaVigna and Gentzkow (2019) define uniform pricing a situation in which the coefficient of variation is below a threshold value of 0.01. Figure 2 shows that uniform pricing is rare in the transportation sector.

Figure 2: Freight Price Dispersion
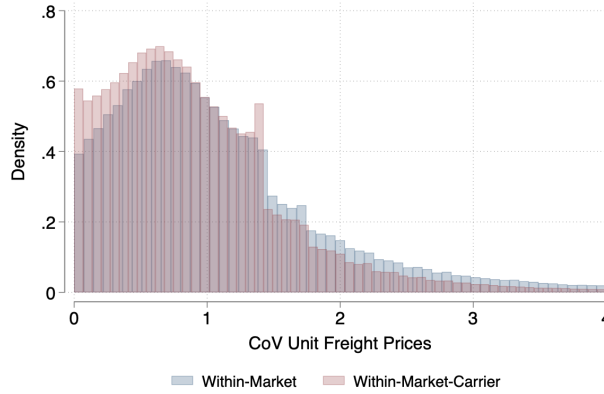


**Notes**: The figure plots the distribution of the coefficient of variation of unit freight prices within a market and within a market-carrier combination (and time). Markets are defined as a mode-origin-sector combination, where modes are sea, air, and road, and sectors are HS4 categories, respectively. Unit freight prices are computed by dividing total freight cost by the quantity transported.

statistical decomposition of unit freight price dispersion:[10]

$$\log p_{ijmt} = \text{FE}_i + \text{FE}_j + \text{FE}_{mt} + \beta \mathbf{X}_{ijmt} + \varepsilon_{ijmt}, \tag{1}$$

where $FE_i$ is an importer fixed effect, $FE_j$ is a carrier fixed effect, $FE_{mt}$ is a market-time fixed effect where a market is a product-origin-mode combination, and $\mathbf{X}_{ijmt}$ represents a set of control variables such as carrier's experience, age of relationship, size of transactions. Panel A of Table 1 reports the variation in prices explained by each component, indicating that firm-level fixed effects cannot capture the full dispersion in bilateral prices. Most of the dispersion is explained by market-time fixed effects (63%) and the match residual (28%). Product and market power heterogeneity across carriers and differences in buyer market power among importers account for a much smaller share of the variance, 3% and 6% respectively. Panel B of Table 1 focuses on the price dispersion within a carrier-market-year, and decomposes it into an importer fixed effect and a match residual component. We find that only 11% of the price dispersion can be explained by heterogeneity across importers. The bulk of the variation (89%) is in fact specific to the carrier-importer relationship within a market-year, consistent with the role of bilateral forces in determining bilateral prices.

As robustness, Table E.1 in Appendix E shows that the decomposition of unit freight price dispersion is quantitatively similar when we measure unit freight prices in terms of unit of value shipped or per kilogram. Table E.2 shows that the decomposition delivers similar results when we consider only the transactions for which the importers explicitly arrange the shipment (according to Incoterms information). Table E.3 instead shows that the carrier-importer match-specific component accounts for 60% of the variation within a market-year when we consider road freight, slightly lower than the 85% and 90% in air and sea freight respectively.

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[10] Refer to Appendix E.3 for additional details.

Table 1: Fixed-effect Decomposition of Freight Price Dispersion

|  | (1) | (2) |
|---|---|---|
| **Panel A - Share of price dispersion explained by:** | | |
| Observables | . | -0.001 |
| Buyer FE | 0.065 | 0.065 |
| Transport Company FE | 0.032 | 0.033 |
| Sector x Time x Origin x Mode | 0.626 | 0.626 |
| Match Residual | 0.277 | 0.277 |
| **Panel B - Within Carrier-Sector-Origin-Time-Mode:** | | |
| Observables | . | 0.000 |
| Buyer FE | 0.118 | 0.117 |
| Match Residual | 0.882 | 0.883 |

**Notes**: The table reports the results of a statistical decomposition exercise based on OLS regressions on the estimating specification in Equation (1). Unit freight prices are computed by dividing total freight cost by the quantity transported. Column (2) includes observable characteristics such as carrier's experience, age of relationship, size of transaction, while Column (1) includes only fixed effects. Markets are defined as a mode-origin-sector combination, where modes are sea, air and road, and sectors are HS4 categories.

**Evidence of Bilateral Bargaining**  We provide reduced-form evidence in line with the presence of importer-carrier bilateral bargaining. In the presence of bilateral bargaining, equilibrium prices reflect at the same time both buyer and seller market power (Alviarez et al., 2023; Antràs and Staiger, 2012). We test whether bilateral prices increase in seller market power, proxied by carrier $j$'s share in importer $i$'s total purchases $s_{ij}$, and decrease in buyer market power, proxied by the importer $i$'s share in carrier $j$'s total sales $x_{ij}$. The economic intuition of the mechanism is as follows: the importance of the importer to the carrier correlates with the markup the carrier can exert. Conversely, the importance of the carrier to the importer correlates with the markdown the carrier can impose. Formally, we consider the following empirical specification:

$$\log p_{ijmt} = \beta_s \log s_{ijmt} + \beta_x \log x_{ijmt} + \beta \mathbf{X_{ijmt}} + FE + \epsilon_{ijmt}, \tag{2}$$

Where $p_{ijmt}$ is the unit freight price paid by importer $i$ to carrier $j$ in market $m$ at time $t$, measured as transport costs per unit shipped; $\mathbf{X_{ijmt}}$ is a set of control variables such as carrier's experience and the age of the bilateral relationship; and $FE$ represents a set of fixed effects. We define a market as an origin-sector-mode combination, where sectors are HS4 categories. We construct instruments for bilateral shares to address their endogeneity with respect to bilateral prices following Alviarez et al. (2023) and exploiting the network structure of the market. More specifically, we instrument the carrier's seller share $s_{ijmt}$ using the (log) sales of $i$'s other carriers to importers other than $i$ in the specific market $m$. For the importer's buyer share $x_{ijmt}$, we use as an instrument the purchases of $j$'s other importers from carriers other than $i$ in the specific market $m$.

In line with economic intuition, Table 2 shows that buyer market power reduces unit freight prices ($\beta_x$), and carrier market power increases unit freight prices ($\beta_s$). The quanti-

Table 2: Prices and Bilateral Concentration

|  | (1) OLS | (2) OLS | (3) OLS | (4) IV |
|---|---|---|---|---|
| Log Carrier Share $s_{ij}$ | 0.089*** | 0.089*** | 0.601*** | 0.505*** |
|  | (0.003) | (0.003) | (0.004) | (0.023) |
| Log Importer Share $x_{ij}$ | -0.262*** | -0.262*** | -0.715*** | -0.608*** |
|  | (0.002) | (0.002) | (0.003) | (0.020) |
| adj. $R^2$ | 0.808 | 0.808 | 0.915 | 0.384 |
| Controls | No | Yes | Yes | Yes |
| $FE_j + FE_i + FE_{mt}$ | Yes | Yes | No | No |
| $FE_{jmt} + FE_{imt}$ | No | No | Yes | Yes |
| $N$ | 2938028 | 2938028 | 2294616 | 2199484 |

**Notes**: The table reports the estimates from the specification in Equation (2). Columns (1) and (2) include carrier, importer, and market fixed effects. Columns (3) and (4) include carrier-market and importer-market fixed effects. Columns (2) to (4) include controls such as carrier's experience and the age of the bilateral relationship. Columns (1) to (3) report OLS estimates; Column (4) reports IV estimates. We exclude all importer-market-time and carrier-market-time singletons from the estimation. Standard errors are clustered at the importer level.

tative effect of buyer and seller market power is similar. In Column (1), which includes importer, carrier, and market fixed effects, a one percent increase in the carrier's share increases unit freight cost by 0.089 pp, and a one percent increase in the importer's share decreases unit freight cost by 0.26 pp. Including additional controls does not impact quantitatively the effects of buyer and seller market power (Column (2)). Including importer-market and carrier-market fixed effects increase the effect of buyer and seller market power to -0.7 pp and 0.6 pp, respectively (specification in Column(3)). Lastly, instrumenting bilateral shares does not substantially impact the effects of buyer and seller market power (Column (4)).

Table E.4 in Appendix E that the qualitative and quantitative results are robust to several alternative specifications. Columns (1) to (3) show that the buyer and seller market power have quantitatively similar effects on unit freight costs independently of whether we consider sea, air, or road freight. Columns (4) and (5) show that quantitatively similar results hold when measuring unit freight prices per kilogram shipped or using the subsample of transactions for which the importers explicitly arrange the shipment (according to Incoterms information).

# 3 Estimating Bargaining Power in Transportation Sector

This section develops and estimates a partial equilibrium theory of bilateral bargaining in the international shipping market. We focus on the determination of shipping prices through a Nash-in-Nash bargaining problem between importers and carriers. The model allows us to estimate key parameters such as the relative bargaining power between importers and carriers, the substitutability across carriers, and the returns to scale of the carriers' production function. The tractability of the framework allows us to embed the key mechanism into a more general model in Section 4.

## 3.1 Theory

The market consists of a finite number of importers, denoted by $i$, and a finite number of carriers, denoted by $j$. We denote the set of carriers to an importer as $J_i$, and the set of importers to a carrier as $Z_j$. We abstract away from endogenous network formation and entry/exit forces, and consider these sets as given.

**Importers** Each importer $i$ produces one good and sells it domestically facing an isoelastic demand function with elasticity $\sigma > 1$. Importers' output is produced by combining an imported intermediate input, $q_{iF}$, with a domestic input, $q_D$, using a constant-return-to-scale production function with unit substitution elasticity between foreign and domestic inputs. Therefore, the share of imported inputs in total cost and the output elasticity of the imported input are constant, and both are denoted by $\gamma$.[11]

As the stylized facts in Section 2 suggest, importers organize the shipment and purchase transportation services. We assume that each unit of imported input requires one unit of transportation services to be delivered to the importers. Thus, the imported input $q_F$ used in production can be written as the output of the following Leontief production function:

$$q_{iF} = \min\{\,\overline{q_{iF}}, t_i\,\}, \tag{3}$$

where $\overline{q_{iF}}$ is the physical imported input, and $t_i$ is the transportation service purchased by the importer.

We assume that importer $i$'s transportation service, $t_i$, represents a composite bundle of carrier-specific varieties. In other words, each importer $i$ purchases a variety of the transportation service from each carrier $j \in J_i$, combining them with a CES technology.[12] Specifically, we write:

$$t_i = \left( \sum_{j \in J_i} t_{ij}^{\frac{\rho-1}{\rho}} \right)^{\frac{\rho}{\rho-1}} \qquad \text{and} \quad \tau_i = \left( \sum_{j \in J_i} \tau_{ij}^{1-\rho} \right)^{\frac{1}{1-\rho}}, \tag{4}$$

where $t_{ij}$ is the quantity of transportation services that importer $i$ purchases from carrier $j$, $\tau_{ij}$ the corresponding bilateral price, and $\rho > 1$ the substitutability across carriers.

It follows that, ultimately, the unit price of imported inputs is $p_{iF} = \overline{p_{iF}} + \tau_i$, where $\overline{p_{iF}}$ is the (possibly $i$-specific) factory-gate price and $\tau_i$ the price index of the bundle of transportation services. We abstract away from any bilateral bargaining between importer and exporter, assuming that the importer is a price taker in the imported input market, taking as given the factory-gate $\overline{p_{iF}}$.

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[11]In other words, $\frac{\partial \log u_i}{\partial \log p_{iF}} = \frac{q_{iF} p_{iF}}{p_i q_i} = \gamma$, where $u_i$ is the marginal cost of importer $i$, while $p_i$ and $q_i$ are the price and the output of importer $i$, respectively.

[12]We show in Appendix B.1 that we can microfound the CES composite bundle of transportation services *via* a discrete choice model in which the importer chooses one single carrier subject to idiosyncratic taste shock distributed according to a Gumbel distribution. This more realistic framework delivers the same implications as our composite bundle of carrier-specific varieties assumption.

**Carriers** On the carrier side, we follow Alviarez et al. (2023) and define the production technology in a parsimonious way. Each carrier sells a unique variety of transportation services to all importers in $Z_j$. We assume that the total costs of production are a function of the total output produced by the carrier, denoted by $t_j$: $TC(t_j) = \zeta_j t_j^{\frac{1}{\theta}}$, where $\zeta_j$ is a constant capturing productivity differences across carriers, and $\theta \in (0,1)$ controls the returns to scale of carriers' production.

Importantly, carriers exhibit an upward-sloped supply curve with marginal cost $c_j$ increasing in quantity, i.e. $\frac{\partial \log c_j}{\partial \log t_j} = \frac{1-\theta}{\theta} > 0$. The presence of decreasing returns to scale in carriers' production guarantees the existence of importers' market power given that the inverse carrier supply elasticity is positive.[13]

**Bargaining over shipment prices** We assume that the bilateral price of transportation services is determined *via* a static, Nash-in-Nash bargaining process (Collard-Wexler et al., 2019; Alviarez et al., 2023). The bilateral price, $\tau_{ij}$, is the outcome of the following maximization taking as given the agreements by all other pairs:

$$\max_{\tau_{ij}} \left(\pi_i(\tau_{ij}) - \widetilde{\pi_{i(-j)}}\right)^{\phi} \left(\pi_j(\tau_{ij}) - \widetilde{\pi_{j(-i)}}\right)^{1-\phi}, \tag{5}$$

where $\phi$ controls the relative bargaining power, and the first (second) term inside parentheses is the gains from trade of importer $i$ (carrier $j$), defined as the payoff from trading with all counterparts in $J_i$ ($Z_j$) minus the payoff from trading with all counterparts except for $j$ ($i$). Specifically, for importer $i$, the gains from trade represent the saving from lower per-unit transportation costs, minus the cost of purchasing services from carrier $j$. Similarly, for carrier $j$, the gains from trade represent the extra revenues from serving importer $i$, net of the additional production costs.

Solving for the first-order condition of the problem in Equation (5), we can write the optimal bilateral price, $\tau_{ij}$, as follows:

$$\tau_{ij} = c_j \mu_{ij} = c_j \left(\omega_{ij} \widehat{\mu_{ij}} + (1 - \omega_{ij}) \overline{\overline{\mu_{ij}}}\right), \tag{6}$$

where $\overline{\overline{\mu_{ij}}}$ is the oligopoly markup, $\widehat{\mu_{ij}}$ is the oligopsony markdown, and $\omega_{ij}$ the effective importers' bargaining power.[14]

The optimal bilateral markup, $\mu_{ij}$, is a weighted average of the markups that arise in the case one side of the market exerts all the bargaining power. Specifically, $\overline{\overline{\mu_{ij}}} = \frac{\epsilon_{ij}}{\epsilon_{ij}-1}$ is the oligopoly markup where $\epsilon_{ij}$ is the perceived demand elasticity of carrier $j$. The elasticity $\epsilon_{ij}$ depends inversely on the share of carrier $j$ in total transportation costs of importer $i$,

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[13]In Appendix B.2, we follow Boehm and Pandalai-Nayar (2022) and show that we can obtain a similar specification for the supply curve and total costs assuming that each carrier has fixed capacity, a reasonable assumption for the international shipping market.

[14]Appendix A.1 provides additional details on the derivations of the key equations, together with analytical expressions for the gains from trade in Equation (5). A richer and more comprehensive analysis of all the special cases embedded in this framework can be found in Alviarez et al. (2023).

$s_{ij} = \frac{\tau_{ij} t_{ij}}{\sum_{z \in J_i} \tau_{iz} t_{iz}}$, so that the carrier charges higher markups the larger is their relevance for the importers' business.[15] Similarly, $\widehat{\mu_{ij}} = \theta \frac{1 - (1 - x_{ij})^{\frac{1}{\theta}}}{x_{ij}}$ is the oligopsony markdown, which depends negatively on the share of total sales of $j$ purchased by importer $i$, $x_{ij} = \frac{t_{ij}}{\sum_{z \in Z_j} t_{zj}}$. In this case, the larger the relevance of an importer in the business of a specific carrier, the higher the markdown they exert.

We interpret the weight $\omega_{ij} = \frac{\overline{\phi} \lambda_{ij}}{1 + \overline{\phi} \lambda_{ij}}$ as the effective importer's bargaining power, that depends positively on the Nash bargaining power parameter, $\overline{\phi} = \frac{\phi}{1 - \phi}$, and negatively on the importers' gains from trade term $\Omega_{ij}$ through the term $\lambda_{ij} = \frac{\sigma - 1}{\epsilon_{ij} - 1} \frac{\gamma s_{i\tau} s_{ij}}{\Omega_{ij}}$.[16] Intuitively, the bilateral price is closer to the oligopolistic case the lower the bargaining power of the importer and/or the larger the gains from trade for the importers.

## 3.2 Estimation

The goal of this section is to estimate the key parameters of our theory: $\phi$, that controls the relative bargaining power between importers and carriers, $\rho$, that governs the substitutability across carriers, and $\theta$, that controls the return to scale of carriers' production function. We use a two-step empirical strategy. We first estimate substitutability across carriers employing a standard IV strategy and the log-log relationship between prices and shares implied by our framework. Then, given the estimated $\rho$, we estimate the remaining parameters leveraging the identification assumption in Alviarez et al. (2023). We use the estimated parameters in Section 4 to perform counterfactual exercises.

**Identification - $\rho$**  The identification of the substitutability across carriers, $\rho$, relies on the demand equation for transportation services. The specification of the model in Equation (4) reveals that, for each importer in a specific market $m$, the observed log of the share of carrier $j$ in total transportation costs of importer $i$, $s_{ijt}^m$, depends linearly on the log of the bilateral price, $\tau_{ijt}^m$:

$$\log s_{ijt}^m = -(\rho - 1)\left(\log \tau_{ijt}^m - \log \tau_{it}^m\right) + \nu_{ijt}^m, \tag{7}$$

where the superscript $m$ refers to a specific market (i.e. product-route pair), $\tau_{it}^m$ is the price index at the importer level, and $\nu_{ijt}^m$ is an idiosyncratic demand shock of importer $i$ for carrier $j$ in market $m$, typically assumed to be i.i.d. across (i, j, m, t) with (conditional) mean zero. Equation (7) translates into the following empirical specification assuming that $\rho$ is constant across all markets and importers:

$$\log s_{ijt}^m = \beta \log \tau_{ijt}^m + \alpha_{it}^m + \nu_{ijt}^m, \tag{8}$$

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[15] It can be shown that the $\epsilon_{ij}$ has the following function form: $\epsilon_{ij} = (1 - s_{ij}) \cdot \rho + s_{ij} \cdot (s_{i\tau} \cdot (1 - \gamma + \sigma \cdot \gamma))$. Further details in Appendix A.1.

[16] We have defined the gains from trade for the importer as $\Omega_{ij} = \left[1 - (1 + s_{i\tau} \Delta \tau)^{\gamma(1-\sigma)}\right]$, where $s_{i\tau} = \frac{\tau_i}{p_{iF}} = \frac{\tau_i}{\overline{p_{iF}} + \tau_i}$ is the share of transportation costs in the price of imported goods, and $\Delta \tau = (1 - s_{ij})^{\frac{1}{1-\rho}} - 1$ is the change in the unit cost of transportation services.

where $\alpha$'s is a set of importer-market-time fixed effects, and $\beta$ is the coefficient of interest. To address the standard endogeneity bias associated with OLS regressions of prices on market shares, we instrument transportation prices using Hausman-type and BLP-type instruments. Specifically, exploiting the presence of multiple markets, we consider the price charged by the same carrier $j$ to *other* importers in *other* markets (Hausman et al., 1994). The instrument is exploiting common carrier-level cost shocks across markets for identification. The key assumption is that importers' demand shocks are not correlated across markets, $cov(\nu_{ijt}^m, \nu_{i'jt}^{m'}) = 0$. This assumption would be violated in the presence of carriers' (unobserved) promotional or advertising campaigns across markets. We do not view this as a compelling scenario, given the nature of the international shipping market.[17]

We estimate the specification in Equation (7) differencing out the importer's price index, $\tau_{it}^m$, which is common across all carriers $j$ for a given importer $i$ in a given market $m$ (Broda and Weinstein, 2006; Feenstra, 1994). Specifically, we take the difference of the bilateral share and price of importer $i$ and carrier $j$ and the bilateral share and price of importer $i$ with a different carrier $j'$ in the same market $m$. Formally, defining $\Delta \log x_{ijj't}^m \equiv \log x_{ijt}^m - \log x_{ij't}^m$, we can rewrite Equation (7) as: $\Delta \log s_{ijj't}^m = -(\rho - 1)\Delta \log \tau_{ijj't}^m + \Delta \nu_{ijj't}^m$. This allows us to estimate the specification in Equation (8) abstracting away from importers-market-time fixed effects. In our preferred specification, we include carrier, time, and market fixed effects, thereby limiting potential endogeneity issues to time-varying pair-specific shocks.

We perform several robustness exercises using different specifications, aggregation levels, and set of instruments. We estimate the main specification in difference including the number of carriers and importers competing in each market as additional instruments. In this case, instruments carry information on the market structure and the identification relies on the standard assumption that the entry of carriers and importers takes place before the realization of the shocks (Berry et al., 1995; Gandhi and Nevo, 2021). In addition, we directly estimate the specification in Equation (8) including importer-market-time and carrier fixed effects, both at yearly and quarterly frequency. Lastly, we estimate a reduced form log-log demand aggregating the market share at the carrier (seller) level and dropping information on the importer (buyer) side (Berry et al., 1995).

**Identification - $\phi$ and $\theta$**   We follow Alviarez et al. (2023) and Dhyne et al. (2022) for the identification of the bargaining power parameter, $\phi$, and the carriers' scale parameter, $\theta$.

From Equation (6), we can write the log bilateral price of transportation services between carrier $j$ and importer $i$ at time $t$ as the sum of the log bilateral markup and the log marginal cost of carrier $j$:

$$\log \tau_{ijt} = \log \mu_{ijt} + \log c_{jt} + \nu_{ijt},$$

where $\nu_{ijt}$ is mean-zero i.i.d. and captures (unobserved) cost differences across the importers of a given carrier driven by, for instance, quality differentiation or customization. Taking the

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[17]Companies operating in international shipping make use of targeted B2B marketing campaigns rather than more common, widespread B2C ads. This is in line with the fact that prices are bargained, as shown in the previous section. **more formal info**

difference between the price carrier $j$ charges to any two distinct importers, $i$ and $k$, we can abstract away from the marginal cost of the carrier and write the following moment condition for every $(i, k, j, t)$:

$$g(\phi, \theta, \Lambda_{jikt}) \equiv \mathbb{E}_u[\nu_{jit} - \nu_{jkt}|\Lambda_{jikt}] \equiv \mathbb{E}_u[\log \tau_{ijt} - \log \mu_{ijt} - (\log \tau_{kjt} - \log \mu_{kjt})|\Lambda_{jikt}] = 0 \quad \forall i, k, j, t, \tag{9}$$

where $\Lambda_{jikt}$ is the relevant information set. The identification of the parameters of interest is guaranteed by the strict monotonicity and invertibility of the moment condition in $\phi$ and $\theta$, and by the non-linearity in the elements of the information set, specifically the bilateral shares $s_{ijt}$ and $x_{ijt}$ (Alviarez et al., 2023).[18]

We estimate the moment condition in Equation (9) using an IV GMM:

$$\min_{\phi, \theta} G(\phi, \theta) Z' W Z G(\phi, \theta)', \tag{10}$$

where $G(\phi, \theta)$ collects all moment condition across all $i - k - j - t$, $W$ the optimal weighting matrix, and $Z$ the vector of instruments. The moment condition implies that the expected difference in carrier $j$'s marginal cost across importers $i$ and $k$ is zero. However, difference in the marginal could be correlated with observables such as bilateral shares, creating endogeneity issues. We rely on Hausman-type instruments in constructing $Z$: we include the number of importers and carriers in each market, and the mean (median) bilateral share in the market excluding the involved pairs $i - j$ and $k - j$ (Hausman et al., 1994; Alviarez et al., 2023). As robustness, we also demean at the market, time, and importer level so that the endogeneity concerns are limited to time-varying pair-specific shocks.

## 3.3 Results

**Data construction** Using a structural model, we estimate the parameters of interest using the whole dataset from 2007 to 2022. We define a market $m$ as an HS2 - country of origin - mode of transportation triplet. We collapse all transaction data at the importer-carrier-market-quarter level and construct the key variables of interest $s_{ijt}^m$, $x_{ijt}^m$, $s_{i\tau}$ and $\tau_{ijt}^m$.[19]

In addition to the cleaning described in Section 2.1, we use the following criteria for sample selection. First, we drop observations with zero shares or unit transport price and trim unit transport price at the 5% level within each route and at the 1% level in the whole sample.

For the structural estimation of $\rho$, we further exclude i) markets with only one carrier operating (no variation in $s_{ijt}^m$); ii) carriers operating in only one market or selling only to

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[18]The same moment condition is not strictly monotonic in $\rho$, thus it does not ensure its identification.

[19]We aggregate all transactions at the importer-market-time level and construct the share of transportation services in the price of imports, $s_{i\tau}$, as $\frac{\sum_{jm} \tau_{ijt}^m t_{ijt}^m}{\sum_{jm} (\overline{p}_{iFt}^m + \tau_{ijt}^m t_{ijt}^m)} = \frac{\sum_{jm} \text{Freight Cost}_{ijt}^m}{\sum_{jm} (FOB_{ijt}^m + \text{Freight Cost}_{ijt}^m)}$.

Table 3: Summary Statistics

|  | Mean | Std. |
|---|---|---|
| Log $\tau_{ijt}^m$ | 0.467 | 1.928 |
| Importer's Share $s_{ijt}^m$ | 0.236 | 0.204 |
| Carrier's Share $x_{ijt}^m$ | 0.072 | 0.178 |
| Transport Share $s_{itm}^\tau$ | 0.114 | 0.115 |
| Number of Carriers per Market | 5.292 | 4.485 |
| Number of Importers per Market | 22.074 | 87.710 |
| Number of Carriers per Importer | 2.988 | 0.845 |
| Number of Importers per Carrier | 15.020 | 41.807 |

**Notes**: The table shows the mean and standard deviation for key variables. $\tau_{ijt}^m$ is the unit freight price paid by importer $i$ to carrier $j$ in market $m$ at time $t$, where unit freight price is computed by dividing total freight cost by the quantity transported; $s_{ijt}^m$ is the share of carrier $j$ on importer $i$'s total imports from market $m$ at time $t$; $x_{ijt}^m$ is the share of importer $i$ in $j$'s total quantity transported in market $m$ at time $t$. $s_{imt}^\tau$ is the share of transportation services in the price of imports at the importer-market-time level. A market is defined as a mode-origin-sector combination, where modes are sea, air and road, and sectors are HS2 categories. Appendix **??** provides additional information on the cleaning process.

one importer because we cannot construct the Hausman-type instrument.[20] When estimating $\rho$ in difference, we also exclude importers that purchase transportation services from only one carrier within a market.

For the structural estimation of $\theta$ and $\phi$, we only keep markets in which at least three carriers operate and importers transacting with at least two carriers to ensure enough variation within each market. Moreover, we exclude carriers transacting with only one importer within each market because the moment condition is not defined.

We drop all importer-carrier-market triplets that imply a carrier's perceived demand elasticity $\epsilon_{ijt}^m$ lower than one, which is inconsistent with our model.[21] Because of this restriction, we approximately drop 5% of the observations. Appendix **??** provides additional information on the cleaning process.

Table 3 reports the summary statistics on our sample. As analysed in Section 2, bilateral prices are highly dispersed, and the concentration is high in both the import market and the transportation market. The average number of importers and carriers across markets is 22 and 5, respectively. Importers and carriers are connected to a limited number of partners, translating into high and dispersed market shares $s_{ijt}^m$ and $x_{ijt}^m$. Lastly, the share of transportation services in the price of imports, $s_{imt}^\tau$, is on average 11%, indicating the quantitative relevance of transportation costs for importers. Table C.1 in Appendix **??** shows that the summary statistics are quantitatively similar across modes of transportation.

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[20]We construct the Hausman instrument as:

$$\text{IV}_{ij}^m = \frac{\sum_{m'} \sum_{i'} \tau_{i'j}^{m'} - \sum_{i'} \tau_{i'j}^m - \sum_{m'} \tau_{ij}^{m'}}{\sum_{m'} \sum_{i'} - \sum_{i'} - \sum_{m'},}$$

[21]See Footnote 15.

Table 4: Estimated $\widehat{\rho}$

|  | (1) OLS | (2) IV | (3) IV | (4) IV |
|---|---|---|---|---|
| $\widehat{\beta}$ | 0.063 (0.004) | -1.562 (0.127) | -2.280 (0.663) | -2.014 (0.449) |
| Implied $\widehat{\rho}$ |  | 2.562 | 3.280 | 3.014 |
| FEs | $-$ | $-$ | $FE_j$ | $FE_j + FE_t + FE_m$ |
| N | 1673962 | 1672850 | 1672825 | 1672268 |

**Notes**: The table reports the estimated price elasticities from the specification in Equation (8) estimated in difference. Column (1) reports the OLS estimate. In all IV specifications (Columns (2) to (4)), the vector of instruments includes the average price charged by the same carrier $j$ to other importers in other markets. Standard errors are clustered at the importer level. Implied $\widehat{\rho}$ reports the implied $\rho$, computed as $\widehat{\rho} = -\widehat{\beta} + 1$.

**Estimates of $\rho$** Table 4 shows the results from the estimation of Equation (7) in difference. As expected, we find that the OLS estimate of the price elasticity is positive, displaying a bias towards zero due to the positive correlation between demand and price shocks (Column (1)). When we instrument prices, the magnitude of the estimated CES elasticity decreases relative to its OLS counterpart to approximately -1.5 (Column (2)), confirming the need for price instruments to correct for the endogeneity bias in this setting. Column (3) and Column (4) further saturate the specification in difference including carrier and carrier, market, and time fixed effects, respectively. The point estimates are not particularly sensitive to the set of fixed effects. For each specification, we report the implied substitutability across carriers, $\widehat{\rho}$. Our preferred specification in Column (4) precisely estimates $\widehat{\rho}$ to be three, indicating a low substitutability across carriers within each market.

Table F.1 in Appendix F shows that the estimated price elasticity and substitutability across carriers is robust across several specifications. Estimating the specification in level with the inclusion of importer-market-time fixed effects delivers a price elasticity of -1.3 and -1.8 at yearly and quarterly frequency, respectively. Similarly, we obtain values of $\widehat{\rho}$ between 2.7 and 3.6 when aggregating at the carrier level.

**Estimates of $\phi$ and $\theta$** Before estimating the two parameters central to the bargaining process, we set the elasticity of substitution across carrier to the estimated value of three, $\rho = 3$, and calibrate the values of the parameters $\sigma$ and $\gamma$ to be 6 and 0.5, respectively.[22]

Figure 3 displays the distribution of the estimated parameters across markets. The average and the median bargaining power of the importers, $\phi$, across markets are 0.72 and 0.83, respectively. The implied average importers' relative bargaining power is $\overline{\phi} = \frac{\phi}{1-\phi} = 2.57$, indicating that, on average, importers enjoy a substantial degree of buyer market power. The

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[22]$\sigma$ and $\gamma$ are calibrated using firm-level data from Chilean manufacturing sectors, as described in the quantitative model in Section 4.

Figure 3: Distribution Estimated Parameters
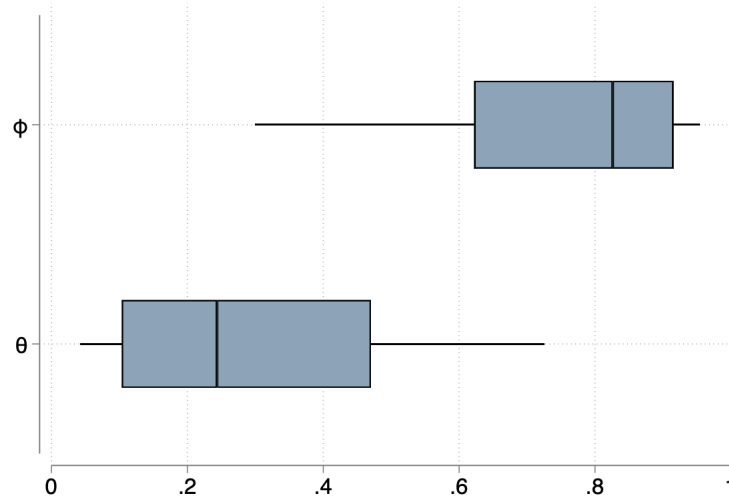


**Notes**: The Figure plots the distribution of the estimated bargaining power parameter $\phi$ (top row) and return to scale parameter $\theta$ (bottom row). The box delimits the interquartile range of the distribution, while the whiskers span from the 10th to the 90th percentiles.

average and the median return to scale of the carriers, $\theta$, across markets are 0.3 and 0.21, respectively, far below one. The implied carriers' supply elasticity is $\frac{\theta}{1-\theta} = 0.43$, indicating that importer market power is sensitive to their size. Figure F.1 in Appendix F displays the distribution of the two parameters depending on the mode of transportation, i.e. distinguishing sea, air, and road freight. The distributions are quantitatively similar across modes, with $\theta$ and $\phi$ being slightly higher and lower in the case of road freight, respectively.[23]

There is a substantial heterogeneity in both bargaining power and return to scale across markets, with an interquartile range of 0.29 and 0.37, respectively. We show that the estimated parameters correlate with observable characteristics of the market in an economically meaningful way. In line with economic intuition, Figure F.3 in Appendix F shows that the bargaining power of the importer, $\phi$, is increasing in the number of carriers in the market, and decreasing in the number of importers in the market. Similarly, the estimated return to scale parameter, $\theta$, an inverse measure of importer market power, is decreasing (increasing) in the number of carriers (importers) in the market. Figure F.4 in Appendix F shows qualitatively similar correlation between the estimated parameters and HHI indices of the bilateral shares $s_{ij}$ and $x_{ij}$, which also capture the relative degree of bargaining power.

# 4    Aggregate Implications

In this section, we embed the bargaining framework developed in Section 3 into a rich general-equilibrium model of importing to quantify the effects that bilateral bargaining in the international shipping market has on the aggregate economy and on the transmission of shocks. Additional details on the derivations are in Appendix A.2.

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[23]Figure F.2 in Appendix F shows the presence of a negative correlation between the estimated importers' bargaining power and the carriers' return to scale across markets, in line with economic intuition.

## 4.1 Theory

**Consumption and Demand** The economy is populated by a unit measure of consumers who supply $L$ units of labor inelastically. They consume a final consumption bundle $C$ over a fixed and exogenous number of domestic products $N$:

$$C = \left( \sum_i^N c_i^{\frac{\sigma-1}{\sigma}} \right)^{\frac{\sigma}{\sigma-1}}, \tag{11}$$

where $\sigma > 1$ is the constant elasticity of substitution across the products in the consumption basket. Consumers maximize their utility subject to a standard budget constraint: $\sum_i^N p_i c_i \leq wL + \sum_i^N \pi_i$, where $w$ is the wage rate and $\pi_i$ are firms' profits. Thus, the demand for each product $i \in N$ is $c_i = p_i^{-\sigma} P^\sigma Y$, where $P$ is the aggregate price index and $Y$ aggregate income.

**Firms and Input Trade** Each product $i \in N$ is produced by a single monopolistically competitive domestic firm combining labor, $l$, and intermediate inputs, $x_i$, using a CRS Cobb-Douglas technology:

$$y_i = \varphi_i l_i^{1-\beta_i} x_i^{\beta_i}, \tag{12}$$

where $\varphi_i$ represents the firm's idiosyncratic productivity. The intermediate input is a combination of domestic and foreign intermediate inputs, $q_D$ and $q_{iF}$, respectively. These are aggregated using a CES technology:

$$x_i = \left( \eta_i q_{iD}^{\frac{\gamma-1}{\gamma}} + (1-\eta_i) q_{iF}^{\frac{\gamma-1}{\gamma}} \right)^{\frac{\gamma}{\gamma-1}}, \tag{13}$$

where $\eta_i > 0$ is the quality of foreign intermediate inputs relative to the domestic one (capturing home bias), and $\gamma > 1$ captures the substitutability between domestic and foreign intermediates.

The firm has access to foreign inputs after paying a fixed cost of $f$ units of domestic labor. We assume that labor can be hired frictionlessly. The presence of fixed costs implies that domestic producers use foreign inputs in their production process only when the unit cost of production decreases enough via the love of variety channels (Halpern et al., 2015; Gopinath and Neiman, 2014; Antras et al., 2017).

We define a roundabout production in the spirit of Caliendo and Parro (2015), assuming that the domestic intermediate input $q_D$ is also produced using the output of all domestic firms as the final consumption good: $q_{iD} = \left( \sum_v^N y_{iv}^{\frac{\sigma-1}{\sigma}} \right)^{\frac{\sigma}{\sigma-1}}$, where $y_{iv}$ is the output of firm $v$ demanded by a firm $i$. Thus, the price of the domestic input $p_D$ is endogenous so that non-importing domestic firms are also affected by changes in the transportation sector via their purchases of intermediate inputs from importers.[24]

Building on our result from Section 3, we assume that domestic producers import foreign

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[24]Notice that all firms aggregate varieties using the same technology, making the domestic input homogeneous across firms and identical to the final consumption good.

intermediate inputs from the rest of the world, purchasing transportation services according to Equations (3) and (4). We assume there exists a unique market for transportation services (i.e. a unique route from the rest of the world to the domestic economy), populated by a finite number of carriers, with the total cost of production increasing in the quantity of services produced, $t_j$: $TC(t_j) = \frac{1}{\zeta_j} t_j^{\frac{1}{\theta}}$, where $\zeta_j$ is a constant capturing productivity differences across carriers, and $\theta \in (0, 1)$ controls the returns to scale of carriers' production. Bilateral prices of transportation services are determined *via* the static, Nash-in-Nash bargaining process in Equation (5).

Under the above assumptions, the firm's profit maximization problem is:

$$\pi_i = \max\{u_i(\tau_i)^{1-\sigma} \times B - wf\mathbf{1}(q_{iF} > 0)\}, \tag{14}$$

where $u_i$ is the unit cost of production for firm $i$, and $B$ is defined as $B \equiv \frac{1}{\sigma}\left(\frac{\sigma}{\sigma-1}\right)^{1-\sigma} P^{\sigma-1} S$, where $S$ is the aggregate spending in the economy.[25]

Formally, the unit cost is given by:

$$u_i = \frac{1}{\varphi_i} w^{1-\beta_i} p_x^{\beta_i} = \frac{1}{\varphi_i} w^{1-\beta_i} \left(\eta_i^\gamma p_{iD}^{1-\gamma} + (1-\eta_i)^\gamma (\alpha_{iq}\overline{p_{iF}} + \alpha_{it}\tau_i)^{1-\gamma}\right)^{\frac{\beta_i}{1-\gamma}}, \tag{15}$$

where $p_x$ is the price index of the intermediate input bundle $x_i$ in Equation (13). The second term is the price of imports, composed by the factory-gate price set by the exporter, $\overline{p_{iF}}$, and the cost of transportation services, $\tau_i = \left(\sum_{j \in J_i} \alpha_{ij} \tau_{ij}^{\frac{1}{\rho-1}}\right)^{\frac{1}{1-\rho}}$.[26]

**General Equilibrium** Equations (11)-(14) above describe firms' optimal decisions. We close the model in general equilibrium, imposing the equilibrium in the labor market and balanced trade between the domestic economy and the rest of the world. Balanced trade requires that aggregate exports equal total imported intermediate inputs:

$$\sum_i^N p_i y_i^{ROW} = \sum_i^N (1 - s_{iD}) m_i,$$

where $m_i$ denotes total intermediate input spending of firm $i$, and $(1 - s_{iD})$ the share of spending on imported inputs.

An equilibrium is defined as a set of (bilateral) prices $\{w, [p_i], [\tau_{ij}]\}$, labor demands for production and fixed costs, demand for services $[t_{ij}]$, production and consumption $\{[y_i], [c_i], [y_i^{ROW}]\}$, and input demands $\{[q_{iD}], [q_{iF}]\}$ such that firms maximize profits, consumers maximize utility,

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[25]Aggregate spending is defined as a function of $L$ and the model's parameters in the following way $S = L^{net}\frac{\sigma}{(1-\gamma)(\sigma-1)}$.

[26]Notice that, relatively to the specification in Equation (4) in Section 3.1, we consider taste heterogeneity across carriers in importers' transportation demand, $\alpha_{ij}$. Moreover, we generalized the Leontief production function in Equation (3) to consider heterogeneity across importers in the mix of transportation services and imports.

trade is balanced, and labor and goods markets clear.[27]

## 4.2 Calibration and Estimation

We now parametrize the model using Chilean customs and micro data. Our calibration and estimation strategy is as follows. Tables 5 and 6 summarize the parameters and the calibrated values or the moments used in the estimation.

**Bargaining and transportation sector.** We use the estimates from Section 3 to parameterize the bargaining process and the transportation sector. We set the elasticity of substitution across carriers to the estimated value of three, $\rho = 3$. We leverage the distribution of parameters estimated across markets in Section 3 and set the bargaining power $\phi$ to 0.8 and the carriers' return to scale $\theta$ to 0.3, respectively, equal to the average estimated parameter across markets. From Table 3, we assume that the number of carriers operating in the transportation sector is equal to 5, i.e. the average number of carriers in a market. Similarly, we set the number of importers equal to the average number of importers across transportation markets, which is approximately 22.

**Domestic Economy.** We use firm-level balance sheet information from the survey of manufacturing industries (ENIA) from 1995 to 2018 to calibrate the parameters defining the domestic economy and the production process ($\sigma$, $\beta$, and $\gamma$).[28] We follow Oberfield and Raval (2021) and identify the elasticity of substitution $\sigma$ from the firms' profit margin, i.e. $\frac{\text{Revenues}_i}{\text{Costs}_i} = \frac{\sigma}{\sigma-1}$. We compute costs as the sum of wage bill, material and electricity expenditure, and user cost of capital. We set the demand elasticity equal 6, close to the median value in the manufacturing sector of 5.9. We then identify the share of material in the production, $\beta$, leveraging the observed factor shares. Given a value for $\sigma$, the observed material spending share allows us to identify $\beta$: $\frac{m_i}{p_i y_i} = \beta \frac{\sigma-1}{\sigma}$. We set $\beta$ to be 0.45, equal to the median material spending share in the manufacturing sector (0.427). Lastly, we identify the substitutability between domestic and imported inputs noting that firm output can be written as:

$$y_i = A \varphi_i l_i^{1-\beta} m_i^\beta s_{iD}^{-\frac{\beta}{\gamma-1}},$$

where $A$ collects all general equilibrium variables, $m_i$ total intermediate input spending of firm $i$, and $s_{iD}$ is the share of spending on domestic inputs. Thus, we leverage the variation in domestic expenditure shares holding material spending fixed to identify $\gamma$ (Blaum et al., 2018; Zhang, 2017). Using standard structural production function estimation techniques as in Olley and Pakes (1992) and Ackerberg et al. (2015), we estimate a value for $\gamma$ of 3.77, and calibrate it to be 4 in our quantitative model. In general, the calibrated values for

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[27]Formally, the labor market clearing condition is: $L = \sum_i (l_i + f\mathbf{1}(q_{iF} > 0))$. Similarly, the good market clearing condition for each firm $i$ is: $y_i = c_i + y_i^{ROW} + \sum_v^N y_{iv}$.

[28]Additional information and details on the dataset ENIA, its cleaning, and the calibration in Appendix D.

the domestic production process are in line with previous estimates and calibrations (Blaum et al., 2018; Alviarez et al., 2023; Halpern et al., 2015).

**Number of firms and fixed costs of importing.** The fixed cost of importing, $f$, is estimated and pinned down by the share of importing firms in the economy. Chilean manufacturing micro data show that 20% of domestic firms are importers. We also assume that the economy is populated by 110 domestic firms (i.e. $N = 110$) so that the number of importers in the transportation sector is consistent with the empirical share of importing firms.

**Firm and carrier productivity, and carrier-import matching shocks.** Four parameters govern the distributions of firms' heterogeneities. Carriers' productivity, $\zeta_j$, is drawn from a Pareto distribution with tail parameter $\kappa_j$. Domestic firm efficiency, $\varphi_i$, is drawn from a log-normal distribution with variance $\sigma_\varphi^2$ and mean $\mu_{\varphi_i}$. Lastly, we assume that the match-specific taste shocks in the transportation sector, $\alpha_{ij}$, are drawn from a normal distribution with unit mean and variance $\sigma_{\alpha_{ij}}^2$. We estimate these parameters by targeting salient features of the empirical distribution of sales and bilateral shares. Specifically, the dispersion in domestic sales and in carriers' size are informative for the efficiency of importers and carriers. A non-zero mean allows us to capture differences in productivity between domestic firms and carriers, which is calibrated targeting the aggregate share of domestic inputs in the economy. We calibrate the process for $\alpha_{ij}$ targeting the average dispersion and the average maximum in $s_{ij}$ across importers, the average dispersion and the average maximum in $x_{ij}$ across carriers, and their correlation.

**Home bias and price of import.** Without loss of generality, we normalize the economy-wide degree of home bias, $\eta$, to 0.5. We target the average share of transportation services in the price of imported goods, $s_{i\tau}$, and the aggregate share of domestic inputs in the economy to calibrate the factory-gate price of imports, $\overline{p_{iF}}$, which we assume to be independent across importers.

We estimate the parameters of the model using simulated method of moments. Given the finite number of firms populating the economy, we generate simulated data from the model and solve for the equilibrium of XX economies for a given set of parameters. We compute the equivalent model moments for each simulated economy, compute the average moment across economies and compare it to the true moment in the data. We choose the optimal model parameter vector, $\Theta = \{\overline{p_{iF}}, \sigma_{\varphi_i}^2, \mu_{\varphi_i}, \kappa_j, \sigma_{\alpha_{ij}}^2, f\}$, to make simulated model moments close to data moments. We estimate the optimal vector of parameters $\widehat{\Theta_{SMM}}$ such that:

$$\widehat{\Theta_{SMM}} = \Theta : \min_{\Theta} \left(m(\tilde{x}|\Theta) - m(\tilde{x})\right) W \left(m(\tilde{x}|\Theta) - m(\tilde{x})\right)', \tag{16}$$

where $m(\tilde{x})$ is the vector of data moments, and $m(\tilde{x}|\Theta)$ is the vector of simulated model moments. In estimating Equation (16), we employ a mix of stochastic optimization routine

Table 5: Calibrated Parameters

**Panel A: Calibrated Parameters**

| | | |
|---|---|---|
| Transportation Sector and Bargaining Process | | |
| $\rho$ | 3 | Estimated from Section 3.3 |
| $\theta$ | 0.3 | Average Estimate across Markets from Section 3.3 |
| $\phi$ | 0.8 | Average Estimate across Markets from Section 3.3 |
| $N_j$ | 5 | Average Number of Carrier per Market |
| $N_i$ | 22 | Average Number of Importers per Market |
| $\alpha_{iq}, \alpha_{it}$ | 1 | Normalized |
| Domestic Economy | | |
| $\beta$ | 0.45 | Median Share of Materials |
| $\gamma$ | 4 | Estimated using Production Function |
| $\sigma$ | 6 | Median Markup |
| $N$ | 110 | Share of Importers |
| $\eta$ | 0.5 | Normalized |

**Notes**: The Table reports the value of the calibrated parameters and the moment from the data used for calibration. We use data from the import transaction data from the Chilean Customs from 2007 to 2019, and the survey of manufacturing industries (ENIA) from 1995 to 2018. Additional information on data and their cleaning in Appendices C and D.

and non-stochastic search algorithm. We use a two-step efficient SMM to compute standard errors.[29]

**Estimated Parameters and Model Fit**    Panels A and B of Table 6 report the estimated values with their standard errors and the data and simulated moments, respectively. The estimation process, constrained by the overidentified and nonlinear nature of the model, demonstrates an overall good fit and precision.

The model reproduces the aggregate domestic share and the average share of transportation costs in import prices, $s_{i\tau}$. The model achieves this with factory-gate price of imports, $\overline{p_F}$, of 0.5756 and carriers being less efficient than domestic firms, on average.[30] Less efficient carriers translate into higher transportation costs, influencing the relative price of domestic and imported goods and, thus, the aggregate domestic share and the average share of transportation costs.

The model reproduces well the dispersion in sales in the domestic economy; however, despite the carriers' efficiency being ten times more volatile than domestic firms' efficiency, the dispersion in in carriers' aggregate shares in the model falls a bit short relative to the data. The volatility of firms' and carriers' efficiency, together with the volatility in idiosyncratic import-carrier match shocks, jointly target the moments on the distribution of buyer and seller shares. The model reproduces fairly successfully the average max and less so their dispersion and correlation, which are a little too low and a too high, respectively.

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[29]Figure in Appendix plots the relationship between the estimated parameters and selected target moments that hold significant importance for identification, supporting identification.

[30]Given our distributional assumptions, the mean and variance of domestic firms' productivity, $\varphi_i$, are equal to $\exp(\mu_{\varphi_i} + 0.5 * \sigma_{\varphi_i}^2) \approx 0.82$ and $(\exp(\sigma_{\varphi_i}^2) - 1)\exp(2\mu_{\varphi_i} + \sigma_{\varphi_i}^2) \approx 0.24$, respectively. Similarly, the mean and variance of carriers' productivity, $\zeta_j$, are equal to the scale parameter (1.5765) and its squared (2.48), respectively, given that the location and the shape parameters of the Generalized Pareto are normalized to zero.

Table 6: Parametrization

| A. Estimated parameters | Estimate | (Std. Error) |
|---|---|---|
| Std of domestic firm productivity, $\sigma_{\varphi_i}$ | 0.5473 | (0.00026) |
| Mean of domestic firm productivity, $\mu_{\varphi_i}$ | -0.3492 | (0.00001) |
| Pareto tail parameter carrier productivity, $\kappa_j$ | 1.5765 | (0.00063) |
| Std idiosyncratic importer-carrier match shock, $\sigma_{\alpha_{ij}}$ | 0.2690 | (0.00021) |
| Factory-gate price of imports, $\overline{p_{iF}}$ | 0.5756 | (0.00045) |
| Fixed cost of importing, $f$ | 0.00013 | (0.00011) |
| **B. Targeted moments** | **Data** | **Model** |
| Std sales share in domestic economy | 0.0302 | 0.0411 |
| Aggregate domestic share | 0.8885 | 0.8983 |
| Correlation$(s_{ij}, x_{ij})$ | 0.1194 | 0.1403 |
| Average Max $x_{ij}$ | 0.5622 | 0.4288 |
| Average Std in $x_{ij}$ across carriers | 0.2064 | 0.0968 |
| Average Max $s_{ij}$ | 0.3984 | 0.3099 |
| Average Std in $s_{ij}$ across importers | 0.1660 | 0.0772 |
| Average $s_{i\tau}$ | 0.1015 | 0.0841 |
| Std aggregate shares across carriers | 0.1964 | 0.0755 |

**Notes**: Panel A reports the parameters estimated using a two-stage SMM and the corresponding standard errors. Panel B reports the moments in the data and in the simulated model. The model moments are generated as the average between XX economies. The moments from the data are computed using data from the import transaction data from the Chilean Customs from 2007 to 2019, and the survey of manufacturing industries (ENIA) from 1995 to 2018. Additional information on data and their cleaning in Appendices C and D.

**On Identification and Untargeted Moments**   TDB

## 4.3   The Aggregate Impact of Endogenous Trade Costs

With the estimated models at hand, we aim to quantitatively study the aggregate implications of dual market power. In particular, we are interested in understanding the importance of the bargaining mechanism introduced for the determination of transportation prices and, thus, aggregate welfare. We then show that the pass-through of trade tariffs is 40% lower in our model compared to a standard model with iceberg trade frictions. Lastly, we show that carbon policies such as the extension of the EU ETS on the shipping market have negligible effects on aggregate welfare.

**Counterfactual Pricing Policies**   Here on conduct and models w/out bargaining, buyer mp, etc. TBD

**Tariff Shocks and GFT**   We study how the presence of endogenous trade costs due to dual market power and bargaining affects the welfare effects of tariff shocks, relative to a standard iceberg trade cost case. We map a rise in ad valorem tariffs $\tau^{tax}$ into the model with an increase in the factory gate price of imported goods, $\overline{p_F}(1 + \tau^{tax})$. We consider three different scenarios: a 10 p.p., 20 p.p., and 30 p.p. ad valorem tariff. We measure aggregate welfare in

terms of changes in the aggregate price index, but results are qualitatively similar using final consumers' real consumption.

Table 7 shows that the presence of endogenous trade costs reduces the welfare costs of tariff shocks relative to the iceberg case. For each scenario, we compare the estimated welfare effects from our model accounting and non accounting for entry/exit of importers (Columns (3) and (2), respectively) to the iceberg trade cost model (Column (1)).[31] A 10 p.p. ad valorem increases the aggregate price index by 0.76 p.p., indicating a small pass-through of tariffs into final consumers' prices due to the relatively small content of imported goods in the final consumption basket. In the presence of endogenous trade costs, the increase in the aggregate price index is almost 50% lower, 0.41 p.p. when we do not account for entry/exit.

Figure 4 shows that the lower welfare costs of tariffs in the presence of endogenous trade costs is driven by a decline in trade costs, that partially offset the rise in the factory-gate price of imported goods. The rise in the price of imports reduces the demand for imported goods and, consequently, the demand for transportation services. Due to the presence of decreasing returns to scale and importer market power, the price of transportation services drops (first and second bar on the left). This partially offsets the increase in the cost of imported goods for domestic importers ($p_{iF} = \overline{p_F}(1 + \tau^{tax}) \uparrow\uparrow + \tau_i \downarrow$), ultimately reducing the cost for final consumers.

The welfare effects of entry/exit of importers are non-linear, and depends on the size of the tariff shock. Table 7 shows that entry/exit of importers further lower the costs of tariffs for a sufficiently low ad valorem tariff. The reason is that some firms stop importing because the tariff increases the price of imported input, reducing the incentives of domestic firms to engage in importing. As a consequence, the remaining importers can benefit from stronger buyer market power because the $x_{ij}$s increase on average. The stronger buyer market power allows the remaining importers to further reduce their cost of transportation, offsetting even further the rise in the price of imports due to the tariff. The non-linearity is driven by the fact that when a domestic firm stops importing, their marginal cost and therefore the price of their good increases, ultimately increasing the aggregate price index. It is now clear why, when the tariff is particularly large, the presence of entry/exit amplifies the negative effect of tariffs on consumers' welfare. A large tariff induces substantial exit in importing, offsetting the reallocation of bargaining power towards remaining firms, and inducing a stronger price increase.

The third bar in Figure 4 also shows that the introduction of ad valorem tariffs allows carriers to charge higher markups. The reason is that the share of transportation costs in the price of imports decreases because of the tariff, $s_{i\tau} = \frac{\tau_i}{\overline{p_F}(1+\tau^{tax})+\tau_i}$, lowering the perceived elasticity of carriers' demand. This effect is however small and only marginally offset the larger decline in transportation prices due to the returns to scale. Lastly, the fourth bar shows that, even when we abstracting away from entry/exit, the median buyer market share $x_{ij}$

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[31]We compute the welfare effects in the iceberg trade cost model as follows: $\frac{\partial \log P}{\partial \log \overline{p_F}} = \sum_i s_i \frac{\partial \log u_i}{\partial \log \overline{p_F}} \equiv \frac{\beta(1-\sum_i s_i s_{iD})}{1-\beta \sum_i s_i s_{iD}}$, where $s_i = \frac{p_i y_i}{\sum_i (p_i y_i)}$.

Table 7: Pass-through of Tariffs

| Tariff Increase by | Iceberg Trade Costs (1) | Bilateral - No Entry/Exit (2) | Bilateral (3) |
|---|---|---|---|
| 10 % | 0.76 | 0.41 | 0.36 |
| 20 % | 1.52 | 0.79 | 0.77 |
| 30 % | 2.28 | 1.12 | 1.16 |

**Notes**: The Table reports the percentage change in the aggregate price index after the introduction of an ad valorem tariff on imported goods. Column (1) reports the results from a model where transportation costs are iceberg. Column (2) reports the results from our model, fixing the number of importers to the original equilibrium. Column (3) reports the results from our model accounting for entry and exit.

Figure 4: Pass-through of Tariffs - Decomposition
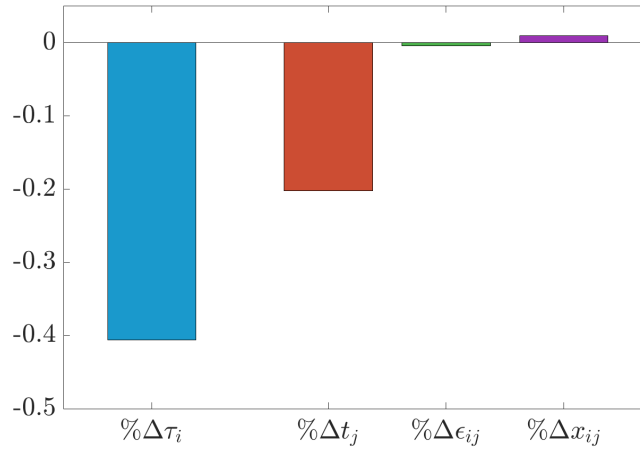


**Notes**: The Figure considers the scenario of a 10 p.p. ad valorem tariff. The blue bar reports the average percentage change in transportation prices. The red bar reports the average percentage change in the scale of the carriers. The green and purple bars report the median change in carriers' perceived demand elasticity and importers' buyer market shares, respectively. All values are computed using the full model, accounting for entry and exit.

increases, indicating a reallocation of market power from larger to smaller importers. The reason is that larger importers, that face lower transportation costs on average, are more exposed to tariff shocks and decrease the demand of imported good relative more. Also in this case, the aggregate effect is small compared to the main channel of the returns to scale.

**Carbon Tax on Transportation**   We leverage the estimated model to assess the potential aggregate welfare effects of the EU carbon tax on international shipping. The shipping sector contributes to the emission of 800 million tonnes of $CO_2$ at the global level, approximately 3% of the total greenhouse emissions.[32] After the creation of a cap-and-trade system in 2005, the EU Emissions Trading System (ETS), the European Commission decided to gradually included international shipping into the EU ETS to reduce its carbon emissions. We map the EU ETS into our model introducing a carbon tax, $\tau^{green}$, that increases the carriers' marginal cost: $\tau_{ij} = \mu_{ij} mc_j (1 + \tau^{green})$. The quantitative exercise applies a €50 tax per ton of $CO_2$, which is

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[32]See Lugovskyy et al. (2023) and https://www.iea.org/energy-system/transport/international-shipping.

mapped to be equivalent to a 5 p.p. increase in the cost of transportation, $\tau^{green} \approx 0.05$.[33] [34] We consider two scenarios: a symmetric one, where all carriers are impacted by the carbon tax; an asymmetric one, where the carbon tax is applied only to one carrier.[35]

Figure 5 shows that the pass-through of the carbon tax is incomplete and the welfare cost for final consumers is negligible. In the symmetric case (bars on the left), transportation prices increase on average by 3.5 p.p., indicating that the carbon tax is passed incompletely by carriers. The reason is that their perceived elasticity (markup) increases (decreases) as the share $s_{i\tau}$ increases, while a minor role can be attributed to the presence of decreasing returns to scale. Despite the increase in transportation prices, the aggregate price index rises only marginally, indicating that the cost for consumer of the carbon tax is negligible.[36] The main reason is that transportation costs account for a small share of importers' marginal cost.

In the asymmetric case, the asymmetry across carriers generates a reallocation of market shares and profits away from the one carrier impacted by the carbon tax. The carrier facing the carbon tax absorbs most of the increase in costs into their margin, increasing their $\tau$ by approximately 1 p.p., much lower than in the symmetric case. The other carriers benefit from the competitive advantage increasing their seller market shares, $s_{ij}$, and exerting stronger market power on importers. This allows them to raise their prices and generate additional profits. Lastly, the fact that the carbon tax is applied to only one carrier explains why the rise in aggregate prices is even smaller relative to the symmetric case.

# 5  Conclusion

This paper examines the role of imperfect competition and bilateral negotiations in the transportation sector and their impacts on international trade. Our analysis provides several key contributions to the literature on international trade and industrial organization.

Using detailed Chilean customs data, we document empirical evidence of high concentration in the transportation sector. We also provide evidence of bilateral negotiations and dual market power between carriers and importers in determining transportation prices. These findings challenge the common assumption in trade literature of perfectly competitive transportation markets and contradict the standard "iceberg" cost assumption prevalent in many trade models.

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[33]We choose a €50 tax per ton of $CO_2$ in line with the EU Commission projections for the 2030s, https://energy.ec.europa.eu/data-and-analysis/energy-modelling/eu-reference-scenario-2020_en, which is lower than the €100 tax per ton of $CO_2$ in Coster et al. (2024) but higher than the €29 tax per ton of $CO_2$ in Shapiro (2016).

[34]We measure the carbon emission of Chilean imports as follows: we first compute the total tonne-km for each mode of transportation using customs and BACI/CEPII data; 2) compute the $CO_2$ emissions generated by applying fuel efficiency coefficients from ECTA/CEFIC (7 gCO2/ton-km for sea freight, 602 gCO2/ton-km for for air freight, and 62 gCO2/ton-km for for road freight); 3) compute the cost of the carbon tax, given a €50 tax per ton of $CO_2$, relative to the total freight costs.

[35]In the asymmetric case, we apply the carbon tax to one carrier at the time and report the average effect across the different iterations.

[36]Notice that we are not accounting for any additional consumers' benefit arising, for instance, from lower carbon emissions.

Figure 5: Carbon Policy
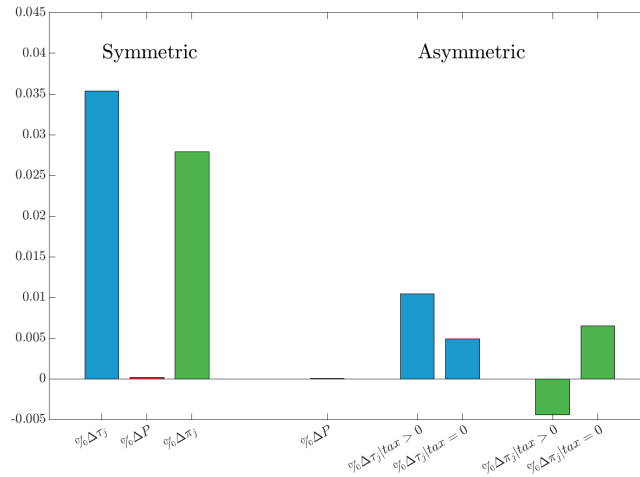


**Notes**: The Figure considers the effect of a carbon tax on carriers' emissions. The set of bars on the left refers to the symmetric case, i.e. all carriers face the carbon tax. The set of bars on the right refers to the asymmetric case, i.e. the carbon tax is applied to one carrier only. In the latter case, we apply the carbon tax to one carrier at the time and report the average effect across the different iterations. The blue bars report the change in the average cost of transportation across carriers; the red bars report the change in the aggregate price index; the green bars report the change in carriers' profits. In the asymmetric case, we distinguish carriers depending on whether they face the carbon tax or not.

We develop a theoretical framework that incorporates bilateral bargaining between carriers and importers, allowing for both seller and buyer market power, and integrate this bilateral bargaining framework into a quantitative trade model. We show that the increase in consumers' prices due to tariffs is 40% lower in our model compared to a standard model with iceberg trade frictions thanks to the presence of decreasing return to scale and importers' market power. We also show that carbon policies such as the extension of the EU ETS on the shipping market have negligible effects on aggregate welfare, but the asymmetry of the policy generates strong reallocation across carriers.

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# A  Derivations and Proofs

## A.1  Derivations of Bilateral Prices

The solution for bilateral transportation prices $\tau_{ij}$ given the framework in Section 3 follows previous work from Alviarez et al. (2023).

**Importer**   Given the assumptions in Section 3, importer $i$'s profits in case of successful negotiations can be written as:

$$\pi_i = (\mu_i - 1)\mu_i^{-\sigma}u_i^{1-\sigma}, \tag{17}$$

where $\mu_i$ is constant and $u_i$ is the marginal cost of importer $i$. It follows that the derivative of $i$'s profits wrt the bilateral transportation price $\tau_{ij}$ is:

$$
\begin{aligned}
\frac{\partial \pi_i}{\partial \tau_{ij}} &= (\mu_i - 1)\mu_i^{-\sigma}(1-\sigma)u_i^{-\sigma}\frac{\partial u_i}{\partial \tau_{ij}}\\
&= (\mu_i - 1)\mu_i^{-\sigma}(1-\sigma)u_i^{-\sigma}\frac{\partial u_i}{\partial p_{iF}}\frac{\partial p_{iF}}{\partial \tau_i}\frac{\partial \tau_i}{\partial \tau_{ij}}\\
&= (\mu_i - 1)\mu_i^{-\sigma}(1-\sigma)u_i^{-\sigma}\gamma\frac{u_i}{p_{iF}}\frac{\tau_{ij}^{-\rho}}{\tau_i^{-\rho}}\\
&= (\mu_i - 1)p_i^{-\sigma}(1-\sigma)\frac{q_{iF}}{q_i}\frac{t_{ij}}{t_i}\\
&= (\mu_i - 1)(1-\sigma)t_{ij}
\end{aligned}
$$

where the last equation is obtained noticing that $t_i = q_{iF}$ given the Leontief production function.

We can derive importer $i$'s profits in case of failed negotiation, $\pi_{i(-j)}$, provided that the cost of a unit of transportation bundle without $j$ is now:

$$\widetilde{\tau}_i = \tau_i(1 - s_{ij})^{\frac{1}{1-\rho}} = \tau_i(1 + \Delta\tau), \tag{18}$$

where $s_{ij} = \frac{\tau_{ij}t_{ij}}{\sum_{z\in J_i}\tau_{iz}t_{iz}}$ is the share of carrier $j$ in total transportation costs of importer $i$. Thus, we can write:

$$\pi_{i(-j)} = (\mu_i - 1)q_i\widetilde{u}_i = (\mu_i - 1)q_iu_i\left(1 + s_{i\tau}\Delta\tau\right)^{\gamma(1-\sigma)},$$

where $s_{i\tau} = \frac{\tau_i}{p_{iF}} = \frac{\tau_i}{p_{iF}+\tau_i}$ is the share of transportation costs in the price of imported goods. It follows immediately that the gains from trade for importer $i$ are:

$$\pi_i(\tau_{ij}) - \pi_{i(-j)} = (\mu_i - 1)q_iu_i\left(1 - (1 + s_{i\tau}\Delta\tau)^{\gamma(1-\sigma)}\right). \tag{19}$$

**Carrier** By the same token, we derive the gains from trade for carrier $j$. The profits of carrier $j$ in case of successful negotiation are:

$$\pi_j(\tau_{ij}) = \tau_{ij}t_{ij} + \sum_{z \neq i \in Z_j} \tau_{zj}t_{zj} - \theta c_j t_j, \tag{20}$$

where $c_j$ is the marginal cost of production given the upward-slope supply curve, and $t_j = \sum_{i \in Z_j} t_{ij}$. It is immediate to show that:

$$\begin{aligned}
\frac{\partial \pi_j(\tau_{ij})}{\partial \tau_{ij}} &= t_{ij} + \tau_{ij}\frac{\partial t_{ij}}{\partial \tau_{ij}} - \theta t_j \frac{\partial c_j}{\partial \tau_{ij}} - \theta c_j \frac{\partial t_j}{\partial \tau_{ij}} \\
&= t_{ij} + \tau_{ij}\frac{\partial t_{ij}}{\partial \tau_{ij}} - c_j \frac{\partial t_{ij}}{\partial \tau_{ij}} \\
&= t_{ij}\left(1 - \epsilon_{ij} - \epsilon_{ij}\frac{c_j}{\tau_{ij}}\right),
\end{aligned}$$

where $\epsilon_{ij} = -\frac{\partial t_{ij}}{\partial \tau_{ij}}\frac{\tau_{ij}}{t_{ij}}$ is the perceived demand elasticity of the carrier. Specifically, given the structure on the importer side,

$$\epsilon_{ij} = (1 - s_{ij}) \cdot \rho + s_{ij} \cdot (s_{i\tau} \cdot (1 - \gamma + \sigma \cdot \gamma)) \tag{21}$$

Moreover, in case of failed negotiations, the profits of carrier $j$ become:

$$\pi_{j(-i)} = \sum_{z \neq i \in Z_j} \tau_{zj}t_{zj} - \theta \widetilde{c}_j \sum_{z \neq i \in Z_j} t_{zj} = \sum_{z \neq i \in Z_j} \tau_{zj}t_{zj} - \theta \widetilde{c}_j t_j(1 - x_{ij}), \tag{22}$$

with $\widetilde{c}_j = c_j\left(1 - x_{ij}\right)^{\frac{1-\theta}{\theta}}$, where $x_{ij} = \frac{t_{ij}}{t_j}$ is the share of total sales of $j$ purchased by importer $i$.

Combining the equations above, we can write the gains from trade for carrier $j$ as:

$$\pi_j(\tau_{ij}) - \pi_{j(-i)} = \tau_{ij}t_{ij} - \theta c_j t_j \left[1 - (1 - x_{ij})^{\frac{1}{\theta}}\right] = t_{ij}(\tau_{ij} - c_j\widehat{\mu_{ij}}), \tag{23}$$

where $\widehat{\mu_{ij}} = \theta\frac{1 - (1 - x_{ij})^{\frac{1}{\theta}}}{x_{ij}}$ is the markup in the oligopsony case.

**Bilateral prices** Given the expressions for the gains from trade above, the FOC for the problem in Equation (5) is:

$$0 = \frac{\partial \pi_j}{\partial \tau_{ij}} + \overline{\phi}\frac{\pi_j - \pi_{j(-i)}}{\pi_i - \pi_{i(-j)}}\frac{\partial \pi_i}{\partial \tau_{ij}},$$

where $\overline{\phi} = \frac{\phi}{1-\phi}$. Substituting the relevant expressions from above, we get:

$$\begin{aligned}
0 &= (1 - \epsilon_{ij} + \epsilon_{ij}\frac{c_j}{\tau_{ij}}) + \overline{\phi}\frac{\tau_{ij} - c_j\widehat{\mu_{ij}}}{q_iu_i\left(1 - (1 + s_{i\tau}\Delta\tau)^{\gamma(1-\sigma)}\right)}(1 - \sigma)t_{ij} \\
&= -1 + \frac{\epsilon_{ij}}{\epsilon_{ij} - 1}\frac{c_j}{\tau_{ij}} - \overline{\phi}\frac{c_j}{\tau_{ij}}\widehat{\mu_{ij}}\frac{1 - \sigma}{\epsilon_{ij} - 1}\frac{\tau_{ij}t_{ij}}{q_iu_i\Omega} + \overline{\phi}\frac{1 - \sigma}{\epsilon_{ij} - 1}\frac{\tau_{ij}t_{ij}}{q_iu_i\Omega}
\end{aligned}$$

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$$= -1 + \overline{\overline{\mu_{ij}}}\frac{c_j}{\tau_{ij}} - \overline{\phi}\lambda_{ij}\widehat{\mu_{ij}}\frac{c_j}{\tau_{ij}} + \overline{\phi}\lambda_{ij}$$

$$\tau_{ij} = c_j\left((1-\omega_{ij})\widehat{\mu_{ij}} + \omega_{ij}\overline{\overline{\mu_{ij}}}\right).$$

which is Equation (6) in the main text, where $\omega_{ij} = \frac{\overline{\phi}\lambda_{ij}}{1+\overline{\phi}\lambda_{ij}}$, $\lambda_{ij} = \frac{\sigma-1}{\epsilon_{ij}-1}\frac{\tau_{ij}t_{ij}}{q_i u_i \Omega}$, $\Omega = \left[1 - (1 + s_{i\tau}\Delta\tau)^{\gamma(1-\sigma)}\right]$, and $\overline{\overline{\mu_{ij}}} = \frac{\epsilon_{ij}}{\epsilon_{ij}-1}$ is the standard markup in case of oligopoly.

## A.2 Derivations of Bilateral Prices in Quantitative Model

Solution for bilateral transportation prices. Given $t_i = \left(\sum_{j\in J_i}\alpha_{ij}^{\frac{1}{\rho}}t_{ij}^{\frac{\rho-1}{\rho}}\right)^{\frac{\rho}{\rho-1}}$ and $\tau_i = \left(\sum_{j\in J_i}\alpha_{ij}\tau_{ij}^{1-\rho}\right)^{\frac{1}{1-\rho}}$:[37]

1. Define failed negotiation bit for importer in nash-bargaining:

$$\pi_{i(-j)} = (\mu_i - 1)\mu_i^{-\sigma}\widetilde{u}_i^{1-\sigma}P^\sigma Y$$

$$\widetilde{u}_i = w^{1-\beta_i}p_x^{\beta_i} = w^{1-\beta_i}\left(\eta_i^\gamma p_D^{1-\gamma} + (1-\eta_i)^\gamma(\alpha_{iq}\bar{p}_F + \alpha_{it}\widetilde{\tau}_i)^{1-\gamma}\right)^{\frac{\beta_i}{1-\gamma}}$$

where $\widetilde{\tau}_i = \tau_i(1-s_{ij})^{\frac{1}{1-\rho}} = \tau_i(1+\Delta\tau)$, with $s_{ij} = \alpha_{ij}\left(\frac{\tau_{ij}}{\tau_i}\right)^{1-\rho}$. We can therefore rewrite $\widetilde{u}_i$ as follows:

$$\widetilde{u}_i = w^{1-\beta_i}p_x^{\beta_i} = w^{1-\beta_i}\left(\eta_i^\gamma p_D^{1-\gamma} + (1-\eta_i)^\gamma(\alpha_{iq}\bar{p}_F + \alpha_{it}\widetilde{\tau}_i)^{1-\gamma}\right)^{\frac{\beta_i}{1-\gamma}}$$

$$= w^{1-\beta_i}\left(\eta_i^\gamma p_D^{1-\gamma} + (1-\eta_i)^\gamma[p_{iF}(1+s_i^\tau\Delta\tau)]^{1-\gamma}\right)^{\frac{\beta_i}{1-\gamma}}$$

$$= w^{1-\beta_i}p_x^{\beta_i}\left(1 + s_i^F[(1+s_i^\tau\Delta\tau)^{1-\gamma} - 1]\right)^{\frac{\beta_i}{1-\gamma}}$$

$$= u_i\left(1 + s_i^F[(1+s_i^\tau\Delta\tau)^{1-\gamma} - 1]\right)^{\frac{\beta_i}{1-\gamma}}$$

where $s_i^\tau = \frac{\alpha_{it}\tau_i}{\alpha_{iq}\bar{p}_F + \alpha_{it}\tau_i}$ is the share of transport cost in the cost of imported inputs; $s_i^F = (1-\eta_i)^\gamma\frac{p_{iF}^{1-\gamma}}{p_x^{1-\gamma}}$ is the share of imported inputs in the mix of intermediate inputs.

2. Gains from trade for importer:

$$\pi_i - \pi_{i(-j)} = (\mu_i - 1)\mu_i^{-\sigma}P^\sigma Y\left(u_i^{1-\sigma} - \widetilde{u}_i^{1-\sigma}\right)$$

$$= (\mu_i - 1)u_i y_i(1-\Omega)$$

where $\Omega = \left(1 + s_i^F[(1+s_i^\tau\Delta\tau)^{1-\gamma} - 1]\right)^{\frac{\beta_i(1-\sigma)}{1-\gamma}}$

3. Moreover:

$$\pi_i = (p_i - u_i)y_i = (\mu_i - 1)\mu_i^{-\sigma}u_i^{1-\sigma}P^\sigma Y.$$

$$\frac{\partial\pi_i}{\partial\tau_{ij}} = (\mu_i - 1)\mu_i^{-\sigma}(1-\sigma)u_i^{-\sigma}P^\sigma Y\frac{\partial u_i}{\partial\tau_{ij}}$$

---

[37]Without loss of generality, we abstract away from importers' idiosyncratic productivity.

$$= (\mu_i - 1)\mu_i^{-\sigma}(1-\sigma)u_i^{-\sigma}P^\sigma Y \frac{\partial u_i}{\partial p_x}\frac{\partial p_x}{\partial p_{iF}}\frac{\partial p_{iF}}{\partial \tau_i}\frac{\partial \tau_i}{\partial \tau_{ij}}$$

$$= (\mu_i - 1)\mu_i^{-\sigma}(1-\sigma)u_i^{-\sigma}P^\sigma Y \beta_i w^{1-\beta_i}p_x^{\beta_i-1}(1-\eta_i)^\gamma \frac{p_{iF}^{-\gamma}}{u_i^{-\gamma}}\alpha_{it}\alpha_{ij}\frac{\tau_{ij}^{-\rho}}{\tau_i^{-\rho}}$$

$$= (\mu_i - 1)p_i^{-\sigma}(1-\sigma)P^\sigma Y \beta_i w^{1-\beta_i}p_x^{\beta_i-1}\frac{q_{iF}}{x}\alpha_{it}\frac{t_{ij}}{t_i}$$

$$= (\mu_i - 1)(1-\sigma)t_{ij}$$

4. Determine bilateral prices: (failed negotiation for carrier is as in Morlacco)

$$\max_{\tau_{ij}} \left(\pi_j - \pi_{j(-i)}\right)^{1-\phi}\left(\pi_i - \pi_{i(-j)}\right)^\phi \tag{24}$$

The foc for the following problem is:

$$0 = \frac{\partial \pi_j}{\partial \tau_{ij}} + \bar\phi \frac{\pi_j - \pi_{j(-i)}}{\pi_i - \pi_{i(-j)}}\frac{\partial \pi_i}{\partial \tau_{ij}}$$

$$= 1 - \epsilon_{ij} + \epsilon_{ij}\frac{c_j}{\tau_{ij}} + \bar\phi\frac{\tau_{ij} - c_j\mu^{OLIGS}}{u_iy_i(1-\Omega)}(1-\sigma)t_{ij}$$

$$= -1 + \frac{\epsilon_{ij}}{\epsilon_{ij}-1}\frac{c_j}{\tau_{ij}} - \bar\phi\lambda_{ij} + \bar\phi\lambda_{ij}\frac{c_j}{\tau_{ij}}\mu^{OLIGS}$$

$$\tau_{ij} = c_j\left(\frac{1}{1+\bar\phi\lambda_{ij}}\mu^{OLIGO} + \frac{\bar\phi\lambda_{ij}}{1+\bar\phi\lambda_{ij}}\mu^{OLIGS}\right)$$

where $\lambda_{ij} = \frac{\sigma-1}{\epsilon_{ij}-1}\frac{1}{1-\Omega}\frac{t_{ij}\tau_{ij}}{u_iy_i} = \frac{\sigma-1}{\epsilon_{ij}-1}\frac{1}{1-\Omega}\beta_i s_i^F s_i^\tau s_{ij}$, where the last ratio is share of variety $j$ in total cost.

## A.3 Derivations of Quantitative Model

**Solution for the aggregate spending S** Let's decompose aggregate spending the following way

$$S = S^C + S^{ROW} + S^X$$

$$= I + \sum_i^N (1 - s_{iD})m_i + \sum_i^N s_{iD}m_i = I + \sum_i^N m_i$$

Recall that

$$\pi_i = (p_i - u_i)y_i = (p_i - \frac{\sigma-1}{\sigma}p_i)y_i$$

$$= \frac{1}{\sigma}p_iy_i = \frac{1}{\sigma}p_ip_i^{-\sigma}P^\sigma Y\frac{P}{P}$$

$$= \frac{1}{\sigma}\left(\frac{p_i}{P}\right)^{1-\sigma}PY = \frac{1}{\sigma}\left(\frac{p_i}{P}\right)^{1-\sigma}S$$

$$\vdots$$

$$\sum_{i=1}^{N} \pi_i = \sum_{i=1}^{N} \frac{1}{\sigma} \left(\frac{p_i}{P}\right)^{1-\sigma} S = \frac{1}{\sigma} S$$

so that we can write the representative consumer spending as

$$I = L + \frac{1}{\sigma} S - \sum_{i=1}^{N} f \mathbf{1}(q_{iF} > 0)\}$$

$$= L^{net} + \frac{1}{\sigma} S$$

where $w = 1$ and $L^{net} = L - \sum_{i=1}^{N} f \mathbf{1}(q_{iF} > 0)\}$. Similarly, for the second element of the aggregate spending decomposition

$$\sum_{i=1}^{N} m_i = \sum_{i=1}^{N} \beta \frac{\sigma - 1}{\sigma} p_i y_i$$

$$= \sum_{i=1}^{N} \beta \frac{\sigma - 1}{\sigma} \left(\frac{p_i}{P}\right)^{1-\sigma} S$$

$$= \beta \frac{\sigma - 1}{\sigma} S$$

So that

$$S = L^{net} + \frac{1}{\sigma} S + \beta \frac{\sigma - 1}{\sigma} S$$

$$= L^{net} \frac{\sigma}{(1 - \beta)(\sigma - 1)}$$

# B  Additional Theoretical Results

## B.1  Micro-foundation for Composite Transportation Bundle

We can microfound our assumption on the existence of a composite bundle of transportation services in Equation (4) from the following discrete choice model.

The importer purchases the transportation services $t_i$ from one carrier. We model the choice of carrier $j$ *via* a discrete choice problem. The indirect utility of importer $i$ from choosing a specific $j$ is:

$$V_{ij} = -\log \tau_{ij} + \frac{1}{1 - \rho} \epsilon_{ij}, \tag{25}$$

where $\tau_{ij}$ is the bilateral price between $i$ and $j$, and $\epsilon_{ij}$ is a stochastic, order-specific taste component. The importer chooses the carrier $j$ that maximizes the indirect utility: $j^\star = \arg\max_{j \in Z_i} V_{ij}$.

We assume that $\epsilon_{ij}$ are distributed according to a Gumbel Extreme-Value type I. Thus, we

can define the probability that importer $i$ chooses carrier $j$ is, $P_{ij}$, as

$$P_{ij} \equiv Pr\left(V_{ij} = \max_{z \in Z_j} V_{iz}\right) = \frac{\tau_{ij}^{1-\rho}}{\sum_{z \in Z_j} \tau_{iz}^{1-\rho}}.$$

We can interpret the probability as the share of $i$'s transportation services purchased from $j$, and define the expected demand of importer $i$ for carrier $j$ transportation services, $t_{ij}$, as

$$t_{ij} = \frac{\tau_{ij}^{1-\rho}}{\sum_{z \in Z_j} \tau_{iz}^{1-\rho}} t_i = \frac{\tau_{ij}^{1-\rho}}{\tau_i^{1-\rho}} t_i \quad \text{with } \tau_i = \left(\sum_{z \in Z_j} \tau_{iz}^{1-\rho}\right)^{\frac{1}{1-\rho}}. \tag{26}$$

Following standard arguments (Anderson et al., 1987), we recognize the demand system generated by Equation (4) in the main text.

## B.2 Carriers' capacity utilization and returns to scale

# C   Customs Data

## C.1   Variable Used

## C.2   Cleaning

### C.2.1   Multi-product transaction

### C.2.2   Shippers cleaning

## C.3   Additional Cleaning for Structural Estimation

Table C.1: Summary Statistics by Mode

|  | Sea | | Air | | Road | |
|---|---|---|---|---|---|---|
|  | Mean | Std. | Mean | Std. | Mean | Std. |
| Log $\tau\_ijt^m$ | -0.811 | 1.301 | 2.269 | 0.980 | -1.742 | 0.678 |
| Importer's Share $s_{ijt}^m$ | 0.233 | 0.194 | 0.240 | 0.217 | 0.245 | 0.186 |
| Carrier's Share $x_{ijt}^m$ | 0.062 | 0.164 | 0.083 | 0.192 | 0.177 | 0.244 |
| Transport Share $s_{imt}^\tau$ | 0.062 | 0.047 | 0.176 | 0.140 | 0.106 | 0.069 |
| Number of Carriers per Market | 5.140 | 4.364 | 5.538 | 4.652 | 4.283 | 3.715 |
| Number of Importers per Market | 23.660 | 103.061 | 21.237 | 69.450 | 8.030 | 14.449 |
| Number of Carriers per Importer | 2.939 | 0.946 | 2.622 | 0.810 | 2.540 | 1.202 |
| Number of Importers per Carrier | 10.239 | 23.298 | 11.072 | 21.415 | 5.975 | 8.403 |

**Notes**: The table shows the mean and standard deviation for key variables by mode of transportation. $\tau_{ijt}^m$ is the unit freight price paid by importer $i$ to carrier $j$ in market $m$ at time $t$, where unit freight price is computed by dividing total freight cost by the quantity transported; $s_{ijt}^m$ is the share of carrier $j$ on importer $i$'s total imports from market $m$ at time $t$; $x_{ijt}^m$ is the share of importer $i$ in $j$'s total quantity transported in market $m$ at time $t$. $s_{imt}^\tau$ is the share of transportation services in the price of imports at the importer-market-time level. A market is defined as a mode-origin-sector combination, where modes are sea, air and road, and sectors are HS2 categories. Appendix **??** provides additional information on the cleaning process.

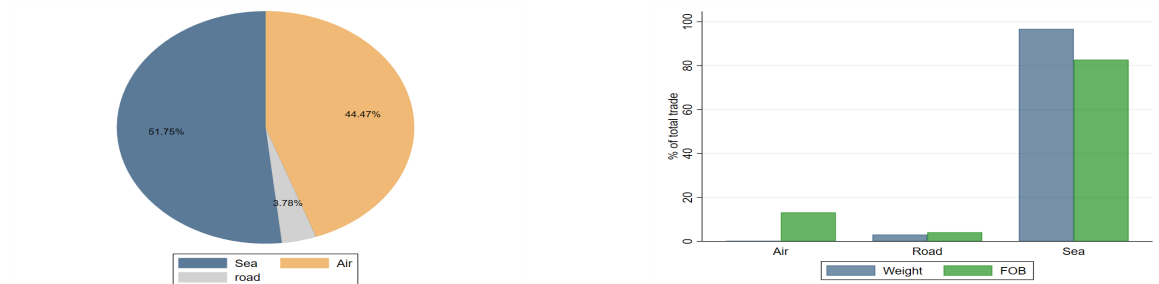# D  Manufacturing Data - ENIA

# E   Additional Data Facts

## E.1   Summary Statistics

In this section we report additional facts on the composition Chilean imports and transportation sector along several dimensions.

### E.1.1   Composition by Mode

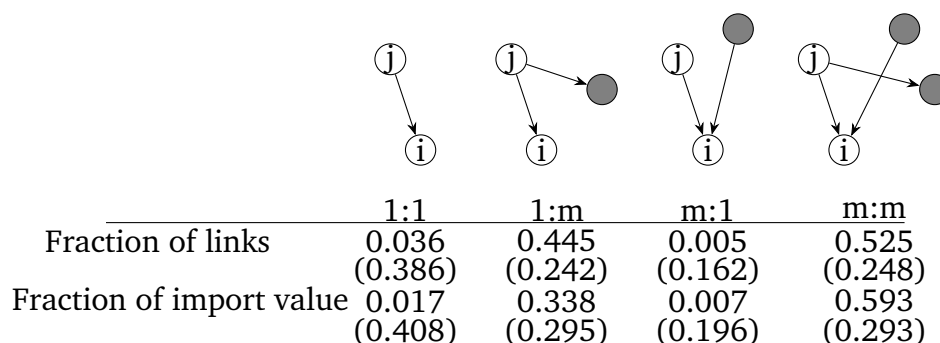Figure E.1: Trade Volumes and Number of Transaction by Mode



**Notes**: The left panel reports the share of total transactions that are conducted via each transport mode. The right panel reports the total value, in navy, and weight, in green, traded by each transport mode. In the customs data, trade via rail is also reported but it represents such a small proportion of total trade (¡1%) that we exclude it from the sample

Figure E.2: Number of modes used



**Notes**: I'll review the grammar and suggest some improvements: This figure reports the share of importers that used one or more transport modes in the sample. The left panel shows the share when the unit of observation is an importer-origin-sector. The right panel shows the same statistic but for a sample in which the unit of observation is an importer-origin.

Figure E.3: Network Structure in International Shipping



| | 1:1 | 1:m | m:1 | m:m |
|---|---|---|---|---|
| Fraction of links | 0.036 | 0.445 | 0.005 | 0.525 |
| | (0.386) | (0.242) | (0.162) | (0.248) |
| Fraction of import value | 0.017 | 0.338 | 0.007 | 0.593 |
| | (0.408) | (0.295) | (0.196) | (0.293) |

**Notes**: This figure illustrate the network structure of the data and the type of relationship that importers and carriers have in our sample. We measure the links both in terms of number of links (top row) and in terms of value traded (bottom row).

### E.1.2 Composition by INCOTERMS

INCOTERMS defines the delivery terms for each transaction. All transactions can be divided into two main groups depending on whether it is the importer's or the exporter's responsibility to arrange the international shipping of the good. In particular, each transaction can be ranked in terms of the importer's responsibility in the delivery process. For instance, the importer plays a fully passive role in the case the agreed term is the so called DDP (Delivered Duty Paid) which places the greatest burden on the exporter. In this case, the exporter agrees in clearing the good through customs at the destination and also to deliver the good at a previously specified location. Thus, when the agreed term is DDP the importer is a spectator in the delivery process. By contrast, under the EXW (Exworks-Factory) the seller has the minimum obligations. Indeed, it is the importer's responsibility to move the good from a designated factory of production to the desired final location. Following standard classification, we group transactions that fall under the category of EXW, FCA (Free Carrier), FOB (Free on Board), and FAS (Free alongside ship) as transactions in which it is the importer's responsibility to arrange the international shipping of the good. By contrast, transactions falling into the terms of CFR (Cost and Freight), CIF (Cost, Insurance and Freight), DDP (Delivered Duty Paid), CPT (Carriage Paid to), DAP (Delivered at Place) are characterized by the fact that it is the seller's responsability to negotiate and pay for the shipping of the good.

Figure E.1: Party Arranging Import Transactions



% Share of Transactions by Delivery Terms
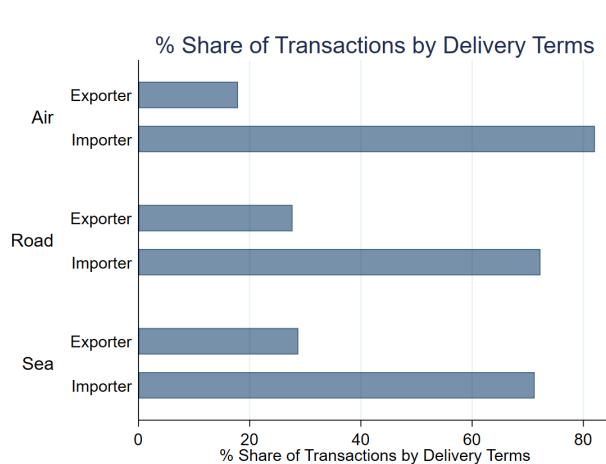
**Notes**: Share of transaction that are arranged by the importer across different transport modes. The party in charge of the transaction is reported in the variable INCOTERMS included in the Chilean custom data.

### E.1.3 Trade Flows Composition

Chilean imports are heterogeneous in terms of product that are brought in from other countries. In Figure E.2 we can see that Chile's imports are spread across different sectors that span natural resource to foods and beverages.

Figure E.2: Import Composition by Sector and Origin
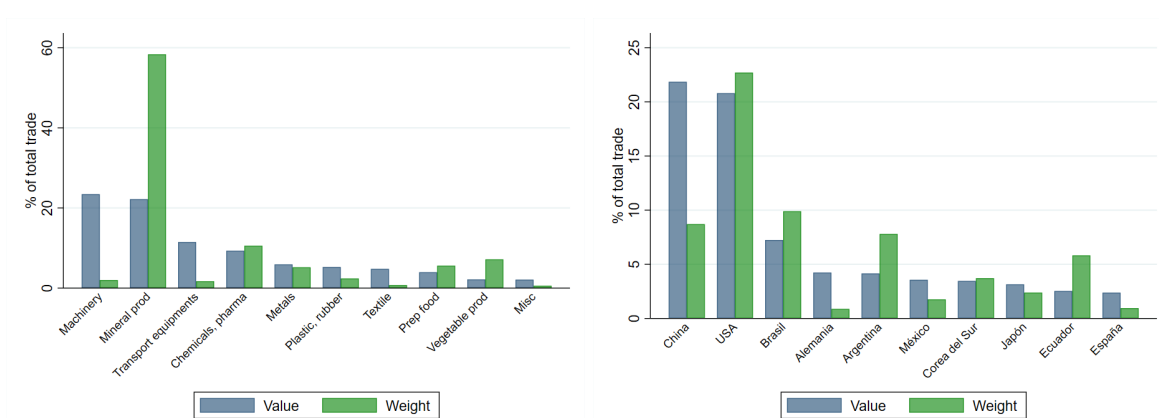


**Notes**: This figure decomposes Chilean imports by sector (left panel) and country of origin (right panel). A sector is defined as one of the 21 sections that compose the more aggregate version of the HS classification. In both figures, the bars are in descending order based on their total value of trade.

Figure E.3: Numbers of Origins by Importer
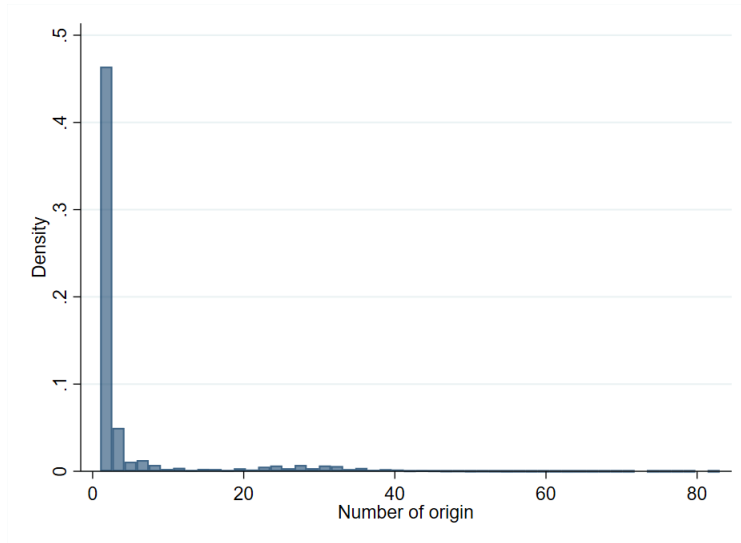


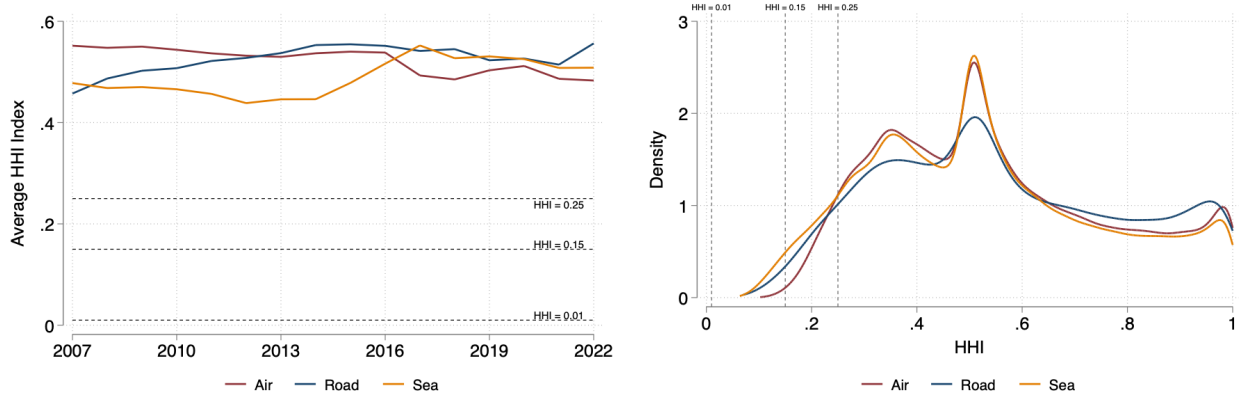**Notes**: This figure reports the distribution of origins per importer.

Figure E.4: Numbers of Products by Importer
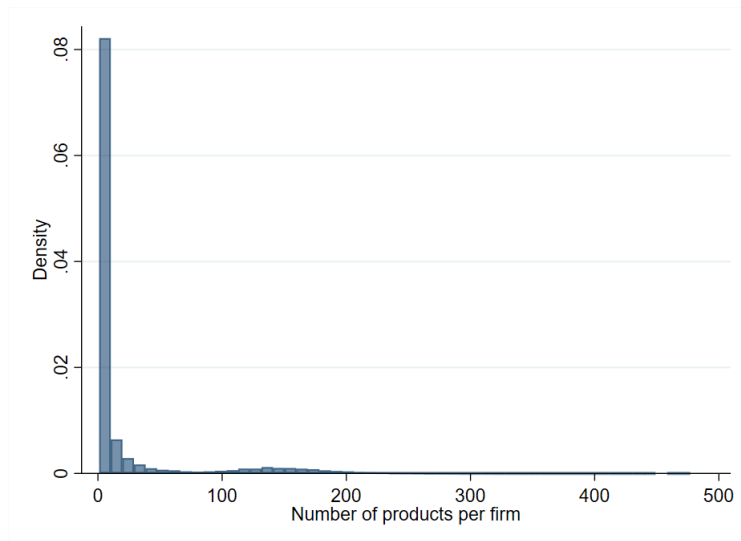


**Notes**: This figure reports the distribution of products per importer.

## E.2   Additional Evidence on Stylized Facts

Figure E.1: Concentration in International Transportation by Mode



**Notes**: The left panel plot the average HHI index across the different markets of the transportation sector over time. Markets are defined as a mode-origin-sector combination, where a sector is defined as a HS4 category. We compute the average distinguishing markets by their mode (sea vs air vs road freight). The left panel plots the distribution of HHI indices across the different markets, distinguishing markets by their mode (sea vs air vs road freight). Carriers' market share are computed in terms of value shipped.

Figure E.2: Concentration in International Transportation - Additional Measures
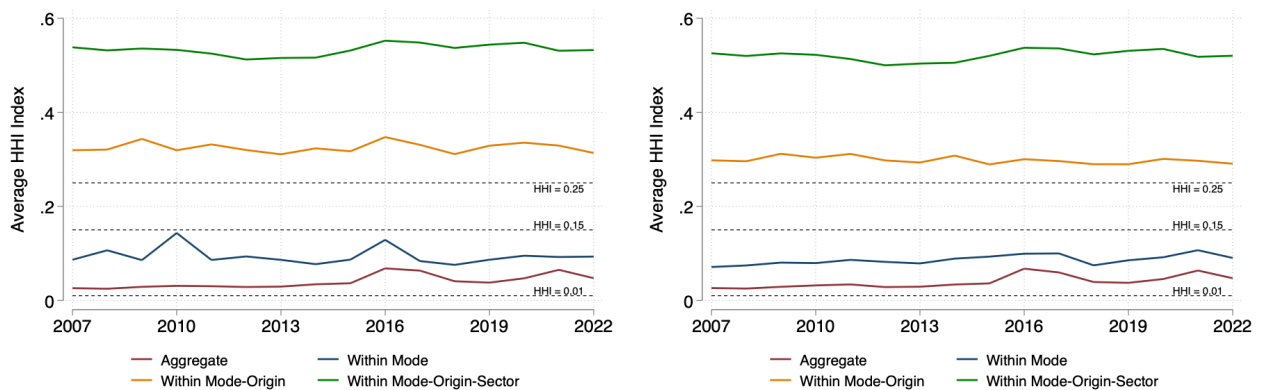


**Notes**: Both panel plot the average HHI index across the different markets of the transportation sector over time. Markets are defined according different levels of granularity. The red line considers a unique aggregate transportation market. The blue line defines markets by their mode of transportation (i.e. sea vs air vs road freight). The orange and green lines defines markets as a combination of mode-origin and mode-origin-sector, respectively. A sector is defined as a HS4 category. The right panel defines carriers' market shares in terms of quantity shipped while the left panel uses kilograms shipped.

Figure E.3: Freight Price Dispersion
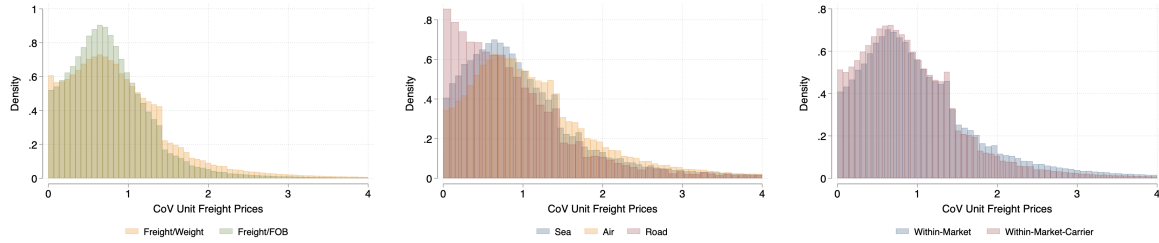


**Notes**: The left panel plots the distribution of the coefficient of variation of unit freight prices within a market (and time). Markets are defined as a mode-origin-sector combination, where modes are sea, air, and road, and sectors are HS4 categories, respectively. Unit freight prices are computed by dividing total freight cost by the weight in kilograms (orange) or by value (green). The center panel plots the distribution of the coefficient of variation of unit freight prices within a market (and time) distinguishing by the mode of transportation (sea, air, and road). Unit freight prices are computed by dividing total freight cost by the quantity transported. The right panel plots the distribution of the coefficient of variation of unit freight prices within a market and within market-carrier pairs (and time) using the subsample of transactions explicitly organized by the importer according to Incoterm information. Unit freight prices are computed by dividing total freight cost by the quantity transported.

Figure E.4: Rejection of Iceberg Trade Cost Assumption
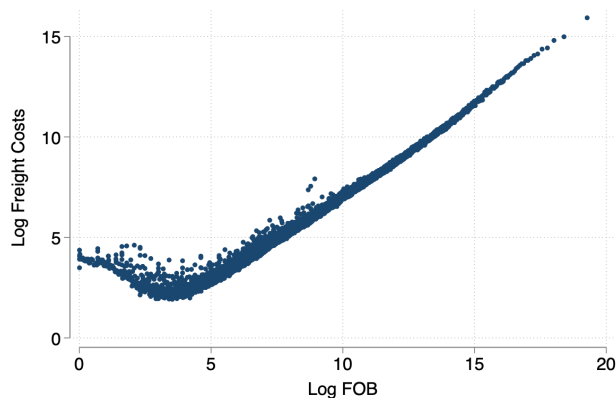


**Notes**: The figure plots the relationship between the log freight costs on the vertical axis and the log of value imported using the whole sample of import transactions.

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Table E.1: Fixed-effect Decomposition of Freight Price Dispersion - Alternative Measures

| | Value | | Weight | |
|---|---|---|---|---|
| | (1) | (2) | (1) | (2) |
| **Panel A - Share of price dispersion explained by:** | | | | |
| Observables | . | 0.069 | . | 0.020 |
| Buyer FE | 0.077 | 0.071 | 0.036 | 0.037 |
| Transport Company FE | 0.022 | 0.024 | 0.046 | 0.047 |
| Sector x Time x Origin x Mode | 0.414 | 0.378 | 0.643 | 0.624 |
| Match Residual | 0.487 | 0.459 | 0.276 | 0.273 |
| **Panel B - Within Carrier-Sector-Origin-Time-Mode:** | | | | |
| Observables | . | 0.090 | . | 0.016 |
| Buyer FE | 0.162 | 0.135 | 0.088 | 0.086 |
| Match Residual | 0.838 | 0.775 | 0.912 | 0.898 |

**Notes**: The table reports the results of a statistical decomposition exercise based on OLS regressions on the estimating specification in Equation (1). Unit freight prices are computed by dividing total freight cost by value (first two columns) or by weight in kilograms (last two columns). Column (2) includes observable characteristics such as carrier's experience, age of relationship, size of transaction, while Column (1) includes only fixed effects. Markets are defined as a mode-origin-sector combination, where modes are sea, air and road, and sectors are HS4 categories.

Table E.2: Fixed-effect Decomposition of Freight Price Dispersion - Incoterms

| | (1) | (2) |
|---|---|---|
| **Panel A - Share of price dispersion explained by:** | | |
| Observables | . | -0.002 |
| Buyer FE | 0.088 | 0.088 |
| Transport Company FE | 0.027 | 0.027 |
| Sector x Time x Origin x Mode | 0.609 | 0.612 |
| Match Residual | 0.275 | 0.275 |
| **Panel B - Within Carrier-Sector-Origin-Time-Mode:** | | |
| Observables | . | 0.001 |
| Buyer FE | 0.154 | 0.153 |
| Match Residual | 0.846 | 0.846 |

**Notes**: The table reports the results of a statistical decomposition exercise based on OLS regressions on the estimating specification in Equation (1). We use the subsample of transactions explicitly organized by the importer according to Incoterm information. Unit freight prices are computed by dividing total freight cost by the quantity transported. Column (2) includes observable characteristics such as carrier's experience, age of relationship, size of transaction, while Column (1) includes only fixed effects. Markets are defined as a mode-origin-sector combination, where modes are sea, air and road, and sectors are HS4 categories.

### Table E.3: Fixed-effect Decomposition of Freight Price Dispersion - by Mode

| | Air (1) | Air (2) | Sea (1) | Sea (2) | Road (1) | Road (2) |
|---|---|---|---|---|---|---|
| **Panel A - Share of price dispersion explained by:** | | | | | | |
| Observables | . | 0.001 | . | 0.009 | . | 0.009 |
| Buyer FE | 0.087 | 0.088 | 0.084 | 0.082 | 0.231 | 0.233 |
| Transport Company FE | 0.007 | 0.007 | 0.034 | 0.033 | 0.033 | 0.037 |
| Sector x Time x Origin x Mode | 0.567 | 0.565 | 0.201 | 0.203 | 0.414 | 0.413 |
| Match Residual | 0.339 | 0.339 | 0.681 | 0.673 | 0.321 | 0.309 |
| | | | | | | |
| **Panel B - Within Carrier-Sector-Origin-Time-Mode:** | | | | | | |
| Observables | . | 0.002 | . | 0.017 | . | 0.042 |
| Buyer FE | 0.155 | 0.156 | 0.101 | 0.098 | 0.398 | 0.380 |
| Match Residual | 0.845 | 0.842 | 0.899 | 0.886 | 0.602 | 0.578 |

**Notes**: The table reports the results of a statistical decomposition exercise based on OLS regressions on the estimating specification in Equation (1). We divide the sample according to the mode of transportation (air, sea, and road). Unit freight prices are computed by dividing total freight cost by the quantity transported. Column (2) includes observable characteristics such as carrier's experience, age of relationship, size of transaction, while Column (1) includes only fixed effects. Markets are defined as a mode-origin-sector combination, where modes are sea, air and road, and sectors are HS4 categories.

### Table E.4: Prices and Bilateral Concentration - Robustness

| | (1) Sea | (2) Air | (3) Road | (4) Weight | (5) Incoterms |
|---|---|---|---|---|---|
| Log Carrier Share $s_{ij}$ | 0.752*** | 0.523*** | 0.585*** | 0.212*** | 0.658*** |
| | (0.004) | (0.005) | (0.035) | (0.005) | (0.005) |
| Log Importer Share $x_{ij}$ | -0.802*** | -0.685*** | -0.607*** | -0.329*** | -0.745*** |
| | (0.003) | (0.004) | (0.035) | (0.004) | (0.004) |
| adj. $R^2$ | 0.887 | 0.749 | 0.958 | 0.845 | 0.910 |
| Controls | Yes | Yes | Yes | Yes | Yes |
| $FE_{jmt} + FE_{imt}$ | Yes | Yes | Yes | Yes | Yes |
| $N$ | 1196245 | 1071462 | 26909 | 2369238 | 1387686 |

**Notes**: The table reports the estimates from the specification in Equation (2) estimated using OLS. All Columns include the additional controls, and carrier-market and importer-market fixed effects. Columns (1), (2), and (3) consider the subsample of sea, air, and road freight, respectively. Column (4) measures unit freight prices per kilogram shipped. Columns (5) reports the estimates using only the subsample of transactions for which the importers explicitly arrange the shipment (according to Incoterms information). We exclude all importer-market-time and carrier-market-time singletons from the estimation. Standard errors are clustered at the importer level.

## E.3   AKM (1999) Decomposition

**Fraction of Links**   The main reference is this paper.

Let $\mathcal{S}$ be a set of $\{1, \ldots, s, \ldots, S\}$ shipper nodes and $\mathcal{M}$ be a set of $\{1, \ldots, m, \ldots, M\}$ of importer nodes. The network structure is a undirected bipartite multi graph with a unbalanced panel structure. The adjacency matrix at a given point in time looks like this

$$A_t = \begin{pmatrix} 0_{S,S} & B_t \\ B_t^{\mathsf{T}} & 0_{M,M} \end{pmatrix},$$

where $B_t$ is a $S \times M$ matrix. Let $e$ be an edge and $E_{(s,m)}$ be the set of edges between $s$ and $m$ (it's a set because it potentially contains repeated links through time) and recall that $E_{(s,m)} = E_{(m,s)}$. Therefore, we define the total (unweighted adjacency matrix) $A$ with elements

$$B_{sm} = \sum_{e \in E_{(s,m)}} \mathbb{1}_e$$

and for the weighted

$$B_{sm}^{w} = \sum_{e \in E_{(s,m)}} w_e$$

where $w_e$ is the import value associated with the transaction that creates the edge between $(s, m)$. Recall that adjacency matrices are symmetric, but $B$ is not

**Decomposition**   The estimated model takes the following form:

$$\log p_{ijopt} = \alpha + \text{FE}_i + \text{FE}_j + \text{FE}_{opt} + \beta \mathbf{X}_{ijopt} + \varepsilon_{ijopt}, \tag{27}$$

where $ijopt$ represent the importer, transportation company, origin country, product (HS4), and quarter, respectively. The dependent variable is the log of the freight value, $\text{FE}_i$ is the time-invariant importer fixed effect, $\text{FE}_j$ is the time-invariant transport company fixed effect, $\mathbf{X}_{ijopt}$ is a vector of control variables, and $\varepsilon_{ijopt}$ is the residual specific to the transport company-importer relationship and the particular product, period, origin, and transport method under consideration. Identification of the transport company and importer fixed effects comes from the variance of freight prices across transport companies and across importers within a product×period×origin×transport-method. The variance of Equation (27) can be decomposed as follow:

$$\begin{aligned} \text{Var}(\log p_{ijopt}) = {} & \text{Cov}(\log p_{ijopt}, \text{FE}_i) + \text{Cov}(\log p_{ijopt}, \text{FE}_j) + \\ & + \text{Cov}(\log p_{ijopt}, \beta \mathbf{X}_{ijopt}) + \text{Cov}(\log p_{ijopt}, \varepsilon_{ijopt}). \end{aligned} \tag{28}$$

# F   Additional Empirical Results

Table F.1: Estimated $\widehat{\rho}$ - Robustness

|  | (1) Additional IV | (2) Level - Yearly | (3) Level - Quarterly | (4) Aggregate | (5) Aggregate |
|---|---|---|---|---|---|
| $\widehat{\beta}$ | -2.101 | -1.306 | -1.759 | -1.700 | -2.606 |
|  | (0.455) | (0.270) | (0.370) | (0.404) | (0.711) |
| Implied $\widehat{\rho}$ | 3.101 | 2.306 | 2.759 | 2.700 | 3.606 |
| FEs | $FE_j + FE_t + FE_m$ | $FE_j + FE_i \times FE_t \times FE_m$ | $FE_j + FE_i \times FE_t \times FE_m$ | $FE_j + FE_t + FE_m$ | $FE_j + FE_t \times FE_m$ |
| N | 1672268 | 2717731 | 2492991 | 502841 | 500097 |

**Notes**: The table reports the estimated price elasticities. Column (1) is estimated in difference and the set of instruments includes the number of carriers and importers in a given market in addition to the average price charged by the same carrier $j$ to other importers in other markets. Column (2) and (3) estimates the specification in level at yearly and quarterly frequency, respectively. The vector of instruments includes only the average price charged by the same carrier $j$ to other importers in other markets. Columns (4) and (5) aggregate data at the carrier-market-year level and estimate the specification in level using the average price charged by the same carrier $j$ in other markets as instrument. Standard errors are clustered at the importer level in Columns (1) to (3), and carrier level in Columns (4) and (5). Implied $\widehat{\rho}$ reports the implied $\rho$, computed as $\widehat{\rho} = -\widehat{\beta} + 1$.

Figure F.1: Distribution Parameters by Mode of Transportation
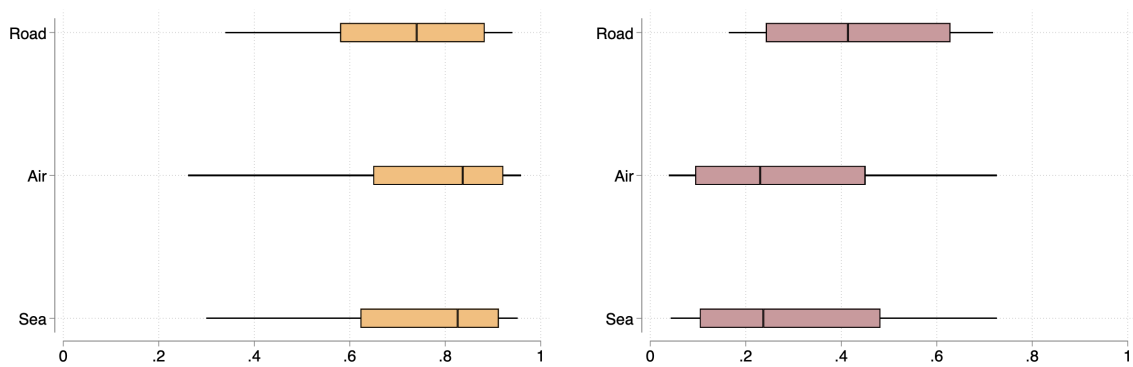


**Notes**: The Figure plots the distribution of the estimated bargaining power parameter $\phi$ (left panel) and return to scale parameter $\theta$ (right panel) by mode of transportation, i.e. distinguishing sea, air, and road markets. The box delimits the interquartile range of the distribution, while the whiskers span from the 10th to the 90th percentiles.

Figure F.2: Correlation $\theta$ and $\phi$ across Markets
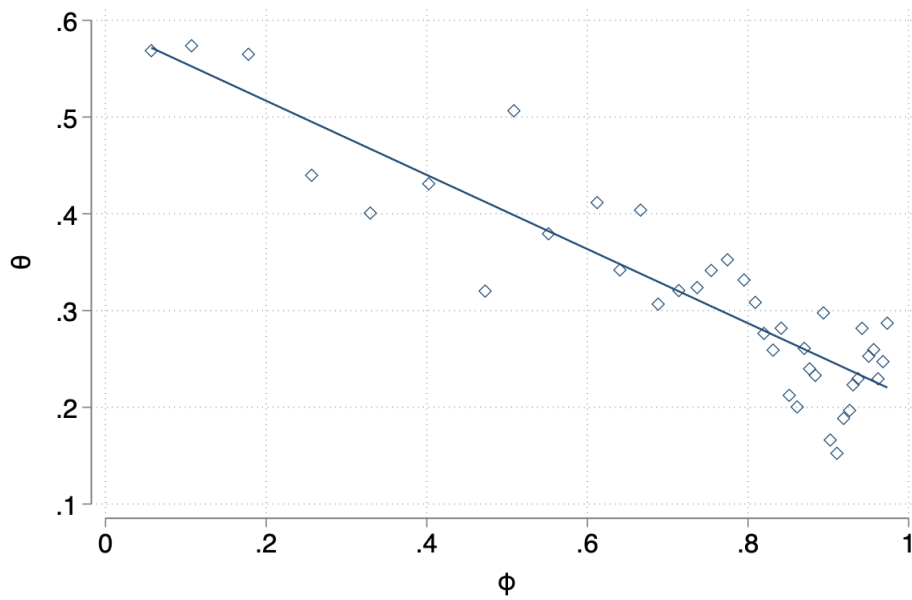


**Notes**: The binscatter plot displays the relationship between the estimated bargaining power (on the horizontal axis) and return to scale (on the vertical axis) parameters across markets.

# Figure F.3: Correlation Estimated Parameters - Number of Firms across Markets
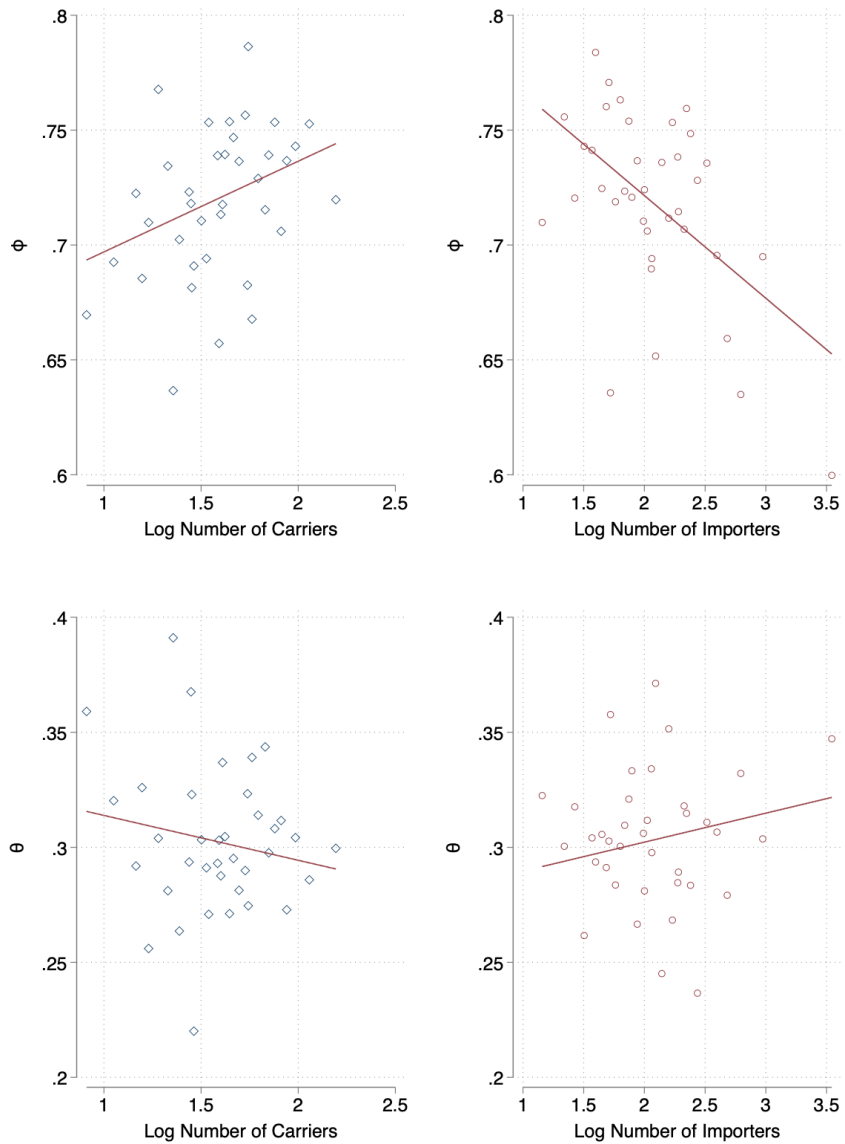


**Notes**: The top (bottom) left panel plots the relationship across markets between the estimated bargaining power (return to scale) parameter and the (log) number of carriers, after residualizing for the (log) number of importers in the market. The top (bottom) right panel plots the relationship across markets between the estimated bargaining power (return to scale) parameter and the (log) number of importers, after residualizing for the (log) number of carriers in the market. In all cases we absorb transport method fixed effects.

## Figure F.4: Correlation Estimated Parameters - HHI across Markets
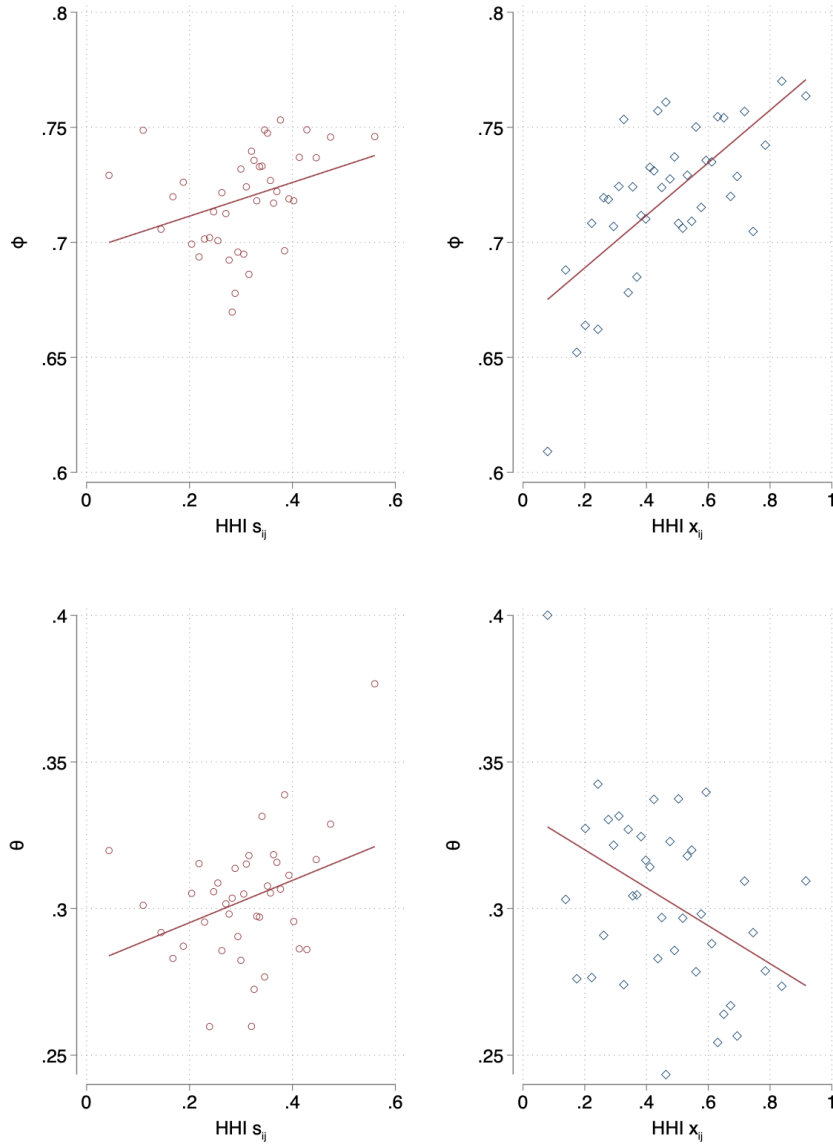


**Notes**: The top (bottom) left panel plots the relationship across markets between the estimated bargaining power (return to scale) parameter and the HHI index of bilateral shares $s_{ij}$ at the market-time level, after residualizing for the HHI index of bilateral shares $x_{ij}$. The top (bottom) right panel plots the relationship across markets between the estimated bargaining power (return to scale) parameter and the HHI index of bilateral shares $x_{ij}$ at the market-time level, after residualizing for the HHI index of bilateral shares $s_{ij}$. In all cases we absorb transport method fixed effects. $s_{ij}$ is the share of carrier $j$ in total transportation costs of importer $i$ (within a market-time pair); $x_{ij}$ is the share of total sales of $j$ purchased by importer $i$ (within a market-time pair).